

We thank the reviewers for their valuable comments. We have made efforts to improve the manuscript accordingly. This document is organized as follows: the referees' comments are in **Bold black**, our responses are in plain black text, and the revisions in the manuscript are shown in blue. The line numbers in this document refer to the updated manuscript.

Anonymous Referee #1

General comments

This manuscript attempts to distinguish contributions from meteorology and emissions reduction to PM_{2.5} trends from 2013 to 2018 in five target regions in China. A multiple linear regression model (MLR) is developed using de-seasonalized (by taking 10-day average of hourly data) and detrended (by subtracting 50-day moving average of 10-day average from 10-day average) PM_{2.5} observations and corresponding five meteorological variables. The coefficients and intercepts obtained for each season and grid are applied to de-seasonalized but not detrended anomalies of meteorological variables (i.e., 50-day moving average minus 6-year average) to calculate PM_{2.5} anomalies attributable to meteorology. Consequently, residual anomalies are attributed to other factors, mainly changes in emissions. The attempt is valuable as the research question, contribution from meteorology to the PM_{2.5} trend, is crucial to East Asian countries. Overall, the results with the MLR is acceptable. I would support publication of this manuscript with minor revision mostly asking clarification.

Specific comments

1) L25 'minor but significant': ambiguous expression. Can you add more explanation?

Thanks. We have rephrased this part to:

The meteorology-corrected PM_{2.5} trends after removal of the MLR meteorological contribution can be viewed as driven by trends in anthropogenic emissions. The mean PM_{2.5} decrease across China is -4.6 μg m⁻³ a⁻¹ in the meteorology-corrected data, 12% weaker than in the original data. The trends in the meteorology-corrected data for the five megacity clusters are: ...

2) L26 'residual anthropogenic trends': anthropogenic emissions?

We have rephrased this sentence to:

The trends in the meteorology-corrected data for the five megacity clusters are: ...

3) Section 2.3: You may consider adding another variable for grid. For now, *i* represents both season and grid which made me difficult to follow at first. Explicit description of $Y_{a,i}(t)$ is needed. It is not clear to me whether the anomaly is $Y_{a,i}(t) = 50\text{-day moving average} - 6\text{-year average at the grid}$ or $Y_{a,i}(t) = 10\text{-day average} - (50\text{-day moving average} - 6\text{-year average})$ at the grid.

Thanks for pointing this out.

$Y_{a,i} = 10\text{-day average} - 6\text{-year average of } 50\text{-day moving average};$

An explanation in brackets (Line 134) is added to explain the way to obtain the PM_{2.5} anomaly $Y_{a,i}$: Consider now the PM_{2.5} anomaly $Y_{a,i}$ for grid square and season i obtained by deseasonalizing but

not detrending the PM_{2.5} data (by removing the 6-year means of the 50-day moving averages), in the same way as for the meteorological variables.

4) Figure S2: How come PM_{2.5} anomalies are greater than deseasonalized and detrended PM_{2.5}? It makes sense if $Y_{a,i}(t)$ is as the second definition as I mentioned above.

PM_{2.5} anomalies ($Y_{a,i}$) can be greater than deseasonalized and detrended PM_{2.5} ($Y_{d,i}$).

$Y_{a,i}$ = 10-day average – 6-year average of 50-day moving average;

$Y_{d,i}$ = 10-day average – 50-day moving average.

From above we can see that trends are not removed from $Y_{a,i}$, and that both trends and seasonal variations are removed from $Y_{d,i}$. Therefore, the difference between PM_{2.5} anomalies and deseasonalized and detrended PM_{2.5} is that PM_{2.5} anomalies contain trend information. This is clarified in the manuscript in Line130 as: “The anomalies calculated in this manner are deseasonalized but not detrended”.

Technical corrections

L131 K. Li et al. (2019): Couldn't find this reference. Did you mean Yi et al. (2019)?

Thanks for pointing this out. We have added this reference in the reference section:

Li, K., Jacob, D. J., Liao, H., Shen, L., Zhang, Q., and Bates, K. H.: Anthropogenic drivers of 2013-2017 trends in summer surface ozone in China, *Proceedings of the National Academy of Sciences*, 116, 422-427, 2019.

We thank the reviewers for their valuable comments. We have made efforts to improve the manuscript accordingly. This document is organized as follows: the referees' comments are in **Bold black**, our responses are in plain black text, and the revisions in the manuscript are shown in blue. The line numbers in this document refer to the updated manuscript.

Anonymous Referee #2

In this study Zhai et al. use a combination of PM_{2.5} observations and multiple linear regression modelling to analyse the trend in PM_{2.5} concentration across mainland China during 2013-2018 and to quantify the meteorological contribution to this trend. Overall the paper is well thought-out and written, and figures are well presented. The topic of the study is interesting and well within the scope of ACP. I recommend publication once the comments below (mostly regarding the processing of the data) have been addressed.

1. Abstract, L25: I suggest specifying that the contribution is “statistically” significant, otherwise the sentence reads a bit odd.

Thanks. To make it clear, we have rephrased this part to:

The meteorology-corrected PM_{2.5} trends after removal of the MLR meteorological contribution can be viewed as driven by trends in anthropogenic emissions. The mean PM_{2.5} decrease across China is -4.6 $\mu\text{g m}^{-3} \text{ a}^{-1}$ in the meteorology-corrected data, 12% weaker than in the original data. The trends in the meteorology-corrected data for the five megacity clusters are: ...

2. Abstract, L26: I think the statement “reduces the uncertainty on the emission-driven trends” needs more explanation in the abstract. Earlier in the abstract you refer to the difficulty of trend attribution because of the meteorologically driven interannual variability in PM_{2.5} concentrations. However, it is not immediately clear what you mean by “uncertainty on the emission-driven trends”. (It is worded more clearly in the conclusions section).

To make it clear, we deleted ‘reduces the uncertainty on the emission-driven trends’, and reworded this part in the abstract as:

The meteorology-corrected PM_{2.5} trends after removal of the MLR meteorological contribution can be viewed as driven by trends in anthropogenic emissions. The mean PM_{2.5} decrease across China is -4.6 $\mu\text{g m}^{-3} \text{ a}^{-1}$ in the meteorology-corrected data, 12% weaker than in the original data. The trends in the meteorology-corrected data for the five megacity clusters are: ...

3. Introduction, L52: Please add an explanation to why the PM_{2.5} concentration is correlated to V850, particularly in the NCP.

Explanation is added in lines 93-94: V850 in particular is a strong predictor of PM_{2.5} wintertime pollution events in the North China Plain, because northerly winds (negative V850) ventilate the region with clean dry air (Cai et al., 2017; Pendergrass et al., 2019).

Added reference:

Cai, W., Li, K., Liao, H., Wang, H., and Wu, L.: Weather conditions conducive to Beijing severe haze more frequent under climate change, *Nature Climate Change*, 7, 257-263, 2017.

4. Introduction, final paragraph: References to a few papers that have identified/quantified recent trends in PM_{2.5} concentrations across China seem to be missing from the introduction (Ma et al., 2019; Silver et al., 2018; Liang et al., 2016).

Added. Thanks.

5. Section 2.1, L88: Can you give any example references here for these previous studies?

Example references (Wang et al., 2014; Cai et al., 2017; Shen et al., 2017; Leung et al., 2018; Song et al., 2019; Zou et al., 2017) are added.

6. Section 2.1, L90: Why was 70% chosen? It seems quite low to me. Please add some justification. Did you do any sensitivity tests changing the threshold to a higher percentage?

This threshold is aimed to include sites that have continuous observations since early 2013 (mainly sites located in the 74 major cities). We have tried to use data from the 74 major cities and obtained identical trend results. I then improved the threshold to 80% and 90% and find that although the number of valid sites in each target region decreased a little bit, the pollutants trends have negligible differences compared with the trends when the '70% threshold' was used.

Added justification in Lines 96-97: [We did sensitivity tests with data coverage thresholds changing from 70% to 90% and obtained similar pollutants trends. To make the most use of available data, 70% is chosen.](#)

7. Section 2.1, L90-91: As above, can you add some justification for the coarse grid chosen? Is this recommended by Tai et al. (2012)? Did you test any other grid resolutions?

It is recommended by Tai et al. (2012) and Shen et al. (2017). I have tried to use 0.5°×0.625° grid resolution. However, finer resolution will result in too few valid grids (grids that have both PM_{2.5} and meteorology observations).

We have reworded the text in the manuscript as: For the MLR model, we further average all data on a 2°×2.5° grid to increase statistical robustness [following Tai et al. \(2012\) and Shen et al. \(2017\)](#).

8. Section 2.1: I see that you removed severe outliers from the observation dataset but what did you do about repeating consecutive values in the dataset (e.g. identified in Rohde and Muller (2015)) and day-to-day repeating sequences of values (e.g. identified by Silver et al., 2018)? If these were not removed, please at least acknowledge that data issues are likely remain in the dataset.

Reply: Thank you for pointing this out. I then checked the impacts of those consecutive repeats on this study. It turned out that these consecutive repeating values have negligible impacts on results in this study. Nevertheless, consecutive repeats identified by Rohde and Muller (2015) and Silver et al. (2018) are unlikely 'realistic' values, and are then removed throughout this study. We removed values from the hourly time series when there are >24 consecutive repeats. The threshold of '>24 consecutive repeats' were chosen by applying a series of thresholds from '>4' to '>24', and we find that different thresholds lead to negligible changes in this study results.

Changes in the manuscript (Line80-83):

At the end of Section 2.1: [There are also occasional consecutive repeats of data that may be caused by faulty instruments or reporting \(Rohde and Muller, 2015; Silver et al., 2018\). Here we removed values](#)

from the hourly time series when there are >24 consecutive repeats. This in whole removed 7.4%, 7.0%, 6.4%, and 6.7% of the PM_{2.5}, SO₂, NO₂, and CO data respectively.

Added reference:

Rohde, R. A., and Muller, R. A.: Air pollution in China: mapping of concentrations and sources, PloS one, 10, e0135749, 2015.

9. Section 2.1: If each year of data and each station were considered separately when applying the 70% data threshold, what does the introduction of more stations/data towards the end of the sampling time period (as more stations come online) do to the trends? I realize the data is averaged over large grid cells, but introduction of many stations later in the time-series (that are not consistent) may impact the trends calculated. Please add some explanation. Did you attempt to calculate the trends based only on data from stations that were online in 2013?

Thanks. This is a good point and we have already considered this in this study. For trend analysis, we only retained data from sites with at least 70% of data coverage for each year from 2013 to 2018. That is, the selected sites must have at least 70% data coverage for each of the 6 years from 2013 to 2018 simultaneously. In this way, we are using consistent sites throughout 2013-2018 for trend analysis. To make it clear, we have made modifications in Lines 96, 106 & 449: [for each of the 6 years from 2013 to 2018.](#)

10. Section 2.2 and 2.3: I am slightly confused by the explanation of the deseasonalizing and detrending process. Perhaps the explanation could be reworded slightly? I have understood it as: the data is de-seasonalized and detrended by taking the 50-day moving average from the 10-day means; whereas the anomalies are deseasonalized (but not detrended) by taking the 6-year mean 50-day moving average from the 10-day means; is this correct?

Yes, this is correct. Explanations in the manuscript are reworded as follows:

P4, Line 112-113: [The deseasonalized and detrended time series are obtained by removing the 50-day moving averages from the 10-day mean time series.](#)

P5, Line 129-130: [We thus apply equation \(1\) to the meteorological anomalies \$X_{a,i,k}\$, obtained by removing the 6-year means of the 50-day moving averages from the 10-day mean time series.](#)

11. Section 3.1: I think this section is really nice and gives some good explanations (with references) for the drivers of the changes in pollutants and/or emissions. However, there is no comparison with previous studies that have calculated trends in PM_{2.5} concentrations over similar time periods (e.g. Ma et al., 2019; Silver et al., 2018; there may be others). Are the calculated trends consistent between studies, despite differences in the data or data processing? I understand that the trends are all calculated over slightly different time periods, but at least a qualitative comparison should be added to the text.

Thanks, the following line is added in P9, Line 168-169: [Trends in China PM_{2.5}, SO₂, and NO₂ presented here are consistent with previous studies \(Silver et al., 2018; Ma et al., 2019\) that cover a shorter time period than 2013-2018.](#)

12. Conclusions: this section is a nice summary of the main points of the paper. However, it would make the results even clearer if the percentage difference from the original trend was

quoted here again as in the abstract and it also might be worth explaining again here what is meant by meteorologically corrected data.

P9, Lines 241-245 are revised as:

We refer to the data series after removal of meteorological variability as the meteorology-corrected data. Thus the 2013-2018 PM_{2.5} decrease for Beijing-Tianjin-Hebei is $-9.3 \pm 1.8 \mu\text{g m}^{-3} \text{ a}^{-1}$ in the original data and is 14% weaker in the meteorology-corrected data ($-8.0 \pm 1.1 \mu\text{g m}^{-3} \text{ a}^{-1}$). For the Sichuan Basin where the meteorological correction is particularly large, the PM_{2.5} decrease is $-6.7 \pm 1.3 \mu\text{g m}^{-3} \text{ a}^{-1}$ in the original data and is reduced by 27% to $-4.9 \pm 0.9 \mu\text{g m}^{-3} \text{ a}^{-1}$ in the meteorology-corrected data.

Explanation of the meteorology-corrected data are also added in the Lines 24-26 and in Lines 137-138.

Additional changes:

1) References updated or newly added:

Cheng, J., Su, J., Cui, T., Li, X., Dong, X., Sun, F., Yang, Y., Tong, D., Zheng, Y., Li, Y., Li, J., Zhang, Q., and He, K.: Dominant role of emission reduction in PM_{2.5} air quality improvement in Beijing during 2013-2017: a model-based decomposition analysis, *Atmos. Chem. Phys.*, 19, 6125-6146, 10.5194/acp-19-6125-2019, 2019.

Song, S., Gao, M., Xu, W., Sun, Y., Worsnop, D. R., Jayne, J. T., Zhang, Y., Zhu, L., Li, M., Zhou, Z., Cheng, C., Lv, Y., Wang, Y., Peng, W., Xu, X., Lin, N., Wang, Y., Wang, S., Munger, J. W., Jacob, D. J., and McElroy, M. B.: Possible heterogeneous chemistry of hydroxymethanesulfonate (HMS) in northern China winter haze, *Atmos. Chem. Phys.*, 19, 1357-1371, 10.5194/acp-19-1357-2019, 2019.

Zou, Y., Wang, Y., Zhang, Y., and Koo, J.-H.: Arctic sea ice, Eurasia snow, and extreme winter haze in China, *Science Advances*, 3, e1602751, 10.1126/sciadv.1602751, 2017.

2) Emission trends in Figure 2 for Pearl River Delta and Sichuan Basin are correctly reversed.

Fine particulate matter (PM_{2.5}) trends in China, 2013-2018: contributions from meteorology

Shixian Zhai^{1,2}, Daniel J. Jacob², Xuan Wang², Lu Shen², Ke Li², Yuzhong Zhang², Ke Gui³, Tianliang Zhao¹, Hong Liao⁴

5 ¹Key Laboratory for Aerosol-Cloud-Precipitation of China Meteorological Administration, Collaborative Innovation Center on Forecast and Evaluation of Meteorological Disasters, School of Atmospheric Physics, Nanjing University of Information Science & Technology, Nanjing 210044, China

²John A. Paulson School of Engineering and Applied Sciences, Harvard University, Cambridge, Massachusetts 02138, USA

³Key Laboratory for Atmospheric Chemistry, Chinese Academy of Meteorological Sciences, CMA, Beijing 100081, China

10 ⁴Jiangsu Key Laboratory of Atmospheric Environment Monitoring and Pollution Control, Collaborative Innovation Center of Atmospheric Environment and Equipment Technology, School of Environmental Science and Engineering, Nanjing University of Information Science & Technology, Nanjing 210044, China

Correspondence to: Shixian Zhai (zhaisx@g.harvard.edu).

Abstract. Fine particulate matter (PM_{2.5}) is a severe air pollution problem in China. Observations of PM_{2.5} have been
15 available since 2013 from a large network operated by the China National Environmental Monitoring Center (CNEMC). The data show a general 30-50% decrease of annual mean PM_{2.5} across China over the 2013-2018 period, averaging $-5.2 \mu\text{g m}^{-3} \text{a}^{-1}$. Trends in the five megacity cluster regions targeted by the government for air quality control are $-9.3 \pm 1.8 \mu\text{g m}^{-3} \text{a}^{-1}$ (\pm 95% confidence interval) for Beijing-Tianjin-Hebei, $-6.1 \pm 1.1 \mu\text{g m}^{-3} \text{a}^{-1}$ for Yangtze River Delta, $-2.7 \pm 0.8 \mu\text{g m}^{-3} \text{a}^{-1}$ for Pearl River Delta, $-6.7 \pm 1.3 \mu\text{g m}^{-3} \text{a}^{-1}$ for Sichuan Basin, and $-6.5 \pm 2.5 \mu\text{g m}^{-3} \text{a}^{-1}$ for Fenwei Plain (Xi'an). Concurrent
20 2013-2018 observations of sulfur dioxide (SO₂) and CO show that the declines in PM_{2.5} are qualitatively consistent with drastic controls of emissions from coal combustion. However, there is also a large meteorologically driven interannual variability of PM_{2.5} that complicates trend attribution. We used a stepwise multiple linear regression (MLR) model to quantify this meteorological contribution to the PM_{2.5} trends across China. The MLR model correlates the 10-day PM_{2.5} anomalies to wind speed, precipitation, relative humidity, temperature, and 850 hPa meridional wind velocity (V850). The meteorology-corrected PM_{2.5} trends after removal of the MLR meteorological contribution can be viewed as driven by trends in anthropogenic emissions. ~~We find that meteorology made a minor but statistically significant contribution to the observed 2013-2018 PM_{2.5} trends across China and that removing this influence reduces the uncertainty on the emission-driven trends.~~ The mean PM_{2.5} decrease across China is $-4.6 \mu\text{g m}^{-3} \text{a}^{-1}$ in the meteorology-corrected data, 12% weaker than in the original data. The ~~residual~~ trends in the meteorology-corrected data for the five megacity clusters ~~attributable to~~

30 ~~changes in anthropogenic emission~~ are $-8.0 \pm 1.1 \mu\text{g m}^{-3} \text{ a}^{-1}$ for Beijing-Tianjin-Hebei (14% weaker than the observed trend),
 $-6.3 \pm 0.9 \mu\text{g m}^{-3} \text{ a}^{-1}$ for Yangtze River Delta (3% stronger), $-2.2 \pm 0.5 \mu\text{g m}^{-3} \text{ a}^{-1}$ for Pearl River Delta (19% weaker), $-4.9 \pm$
 $0.9 \mu\text{g m}^{-3} \text{ a}^{-1}$ for Sichuan Basin (27% weaker), and ~~$-5.04-9 \pm 1.9 \mu\text{g m}^{-3} \text{ a}^{-1}$~~ for Fenwei Plain (Xi'an; ~~235%~~ weaker). 2015-
2017 observations of flattening PM_{2.5} in the Pearl River Delta₇ and increases₅ in the Fenwei Plain₇ can be attributed to
meteorology rather than to relaxation of emission controls.

35 1. Introduction

PM_{2.5} (particulate matter with aerodynamic diameter less than 2.5 μm) is a severe air pollution problem in China, responsible
for 1.1 million excess deaths in 2015 (Cohen et al., 2017). The Chinese government introduced in 2013 the Action Plan on
the Prevention and Control of Air Pollution (Chinese State Council, 2013a), called Clean Air Action for short, to
aggressively control anthropogenic emissions. Starting that year, PM_{2.5} data from a nationwide monitoring network of about
40 1,000 sites~~also~~ became available from the China National Environmental Monitoring Center (CNEMC) of the Ministry of
Ecology and Environment of China (MEEC). These data show 30-40% decreases of PM_{2.5} across eastern China over the
2013-2017 period (Chinese State Council, 2018a; X. Zhang et al., 2019). However, interpretation of these trends in terms of
emission controls may be complicated by interannual variability and trends in meteorology (R. Zhang et al., 2014; Y. Wang
et al., 2014; Zhu et al., 2012; Jia et al., 2015; ~~K~~-Li et al., 2018; Yang et al., 2018; Yang et al., 2016; [Liang et al., 2016](#); Cheng
45 et al., 2019~~8~~; Chen et al., 2019; [Silver et al., 2018](#)). Here we use a stepwise multi-linear regression (MLR) model to separate
the effects of meteorological variability and emission controls on the 2013-2018 trends in PM_{2.5} across China.

Meteorology drives large day-to-day, seasonal, and interannual variations in PM_{2.5} in China by affecting transport,
scavenging, emissions, and chemical production (Y. Wang et al., 2014; Leung et al., 2018; Tai et al., 2012; [Zou et al., 2017](#)).
The relationships between PM_{2.5} and meteorological variables are complex and differ by region and time of year (Shen et al.,
50 2017). For example, wintertime PM_{2.5} pollution events in central and eastern China are associated with low wind speed and
high relative humidity (RH) ([Y](#). Wang et al., 2014; R. Zhang et al., 2014; Pendergrass et al., 2019; Moch et al., 2018; Song et
al., 2019). On the other hand, high wind speeds in northern China in spring and summer promote dust emission (Lyu et al.,
2017; X. Wang et al., 2004). Precipitation scavenging is a major factor driving PM_{2.5} variability in southern and coastal
China (Chen et al., 2018; Leung et al., 2018).

55 Anthropogenic emissions of PM_{2.5} and its precursors including sulfur dioxide (SO₂), nitrogen oxides (NO_x), ammonia (NH₃),
and nonmethane volatile organic compounds (NMVOCs) have undergone large changes in China over the past decades.
Rapid growth in emissions from 1980 to 2006 led to a general increase in PM_{2.5} over China, as demonstrated by visibility

data (Che et al., 2007; Han et al., 2016; Wang and Chen, 2016; Fu et al., 2014; X. Zhang et al., 2012) and since 1999 by satellite aerosol optical depth (AOD) data (Ma et al., 2016; Lin et al., 2018; Zhao et al., 2017). SO₂ emissions peaked in 2006/2007, NO_x emissions peaked in 2011, and NH₃ emissions peaked around 1996, as estimated from emission inventories (Zhao et al., 2017; J. Wang et al., 2017b; Xia et al., 2016; F. Liu et al., 2016a; Lu et al., 2010; Xu et al., 2016; Kang et al., 2016) and observed from satellites (Xia et al., 2016; F. Liu et al., 2016a; de Foy et al., 2016; van der A et al., 2017). SO₂ and NO_x emissions have declined since their peaks, whereas NH₃ emissions have remained relatively stable (Zhao et al., 2017). The onset of emission controls led to slight decreases in PM_{2.5} over the 2006-2012 period as indicated by satellite AOD data (Ma et al., 2016; Lin et al., 2018; Zhao et al., 2017; Ma et al., 2019) and surface observations (Tao et al., 2017; J. Wang et al., 2017). The Clean Air Action greatly increased the scope of emission controls. The Multi-resolution Emission Inventory for China (Zheng et al., 2018) (MEIC, <http://www.meicmodel.org>) estimates nationwide emission decreases over the 2013-2017 period of 59% for SO₂, 33% for primary PM_{2.5}, 21% for NO_x, and 3% for NH₃, with NMVOCs increasing by 2%. Continued reductions in emissions are required and implemented in 2018 (Chinese State Council, 2018b). Quantifying the response of PM_{2.5} to these rapid emission changes by resolving the effect of meteorological variability is an important question for measuring the success of the Clean Air Action.

2. Data and methods

2.1. Observations

We use 2013-2018 hourly data for surface air PM_{2.5} together with SO₂, nitrogen dioxide (NO₂), and carbon monoxide (CO) concentrations from the CNEMC network (<http://106.37.208.233:20035/>). The network started in January 2013 with 496 sites in 74 major cities across the country (Chinese State Council, 2013b), growing to ~1540097 sites in 454 cities by 2018. PM_{2.5} mass concentrations are measured using the micro oscillating balance method and/or the β absorption method (MEP, 2012; Zhang and Cao, 2015). SO₂, NO₂, and CO concentrations are measured at the same sites as PM_{2.5}. NO₂ concentrations are measured by the molybdenum converter method known to have positive interferences from NO₂ oxidation products (Dunlea et al., 2007). SO₂ and CO are respectively measured using ultraviolet fluorescence and infrared absorption (MEP, 2012; Zhang and Cao, 2015). We applied quality control to the hourly CNEMC data following Barrero et al. (2015) to exclude severe outliers. In addition, There are also occasional consecutive repeats of data that may be caused by stuck faulty instruments in the CNEMC data are identified by or reporting (-Rohde and Muller, -(2015;) and Silver et al., -(2018). Here we removed values from the hourly time series when there are >24 consecutive repeats. This in whole removed 76.47%, 75.07%, 65.47%, and 65.79% of the PM_{2.5}, SO₂, NO₂, and CO data respectively.

We correlated these air quality observations with meteorological observations from 839 stations distributed across China (Figure S1) and compiled in the Surface Daily Climate Dataset (V3.0) released by the China National Meteorological Information Center (CNMIC; <http://data.cma.cn/>). These include data for wind speed (WDS), precipitation (PRECIP), relative humidity (RH), and temperature (TEM). We also used the 850-hPa meridional wind velocity (V850) from the MERRA-2 reanalysis produced at $0.5^{\circ} \times 0.625^{\circ}$ horizontal resolution by the NASA Global Modeling and Assimilation Office (<https://gmao.gsfc.nasa.gov/reanalysis/MERRA-2>). We choose these meteorological variables for their strong correlations with $PM_{2.5}$ identified in previous studies ([Y. Wang et al., 2014](#); [Cai et al., 2017](#); [Shen et al., 2017](#); [Leung et al., 2018](#); [Song et al., 2019](#); [Zou et al., 2017](#)). V850 in particular is a strong predictor of $PM_{2.5}$ wintertime pollution events in the North China Plain, because northerly winds (negative V850) ventilate the region with clean dry air (Cai et al., 2017; Pendergrass et al., 2019).

All data in this work are averaged over 10 days (10-day time resolution). Trend analyses use only those sites with at least 70% data coverage for each of the 6 years from of the 2013- to 2018- simultaneously period. We did sensitivity tests with data coverage thresholds changing from 70% to 90% and obtained similar pollutants trends. To make the most use of available data, 70% is chosen. For the MLR model, we further average all data on a $2^{\circ} \times 2.5^{\circ}$ grid to increase statistical robustness following (Tai et al. (2012) and Shen et al. (2017).

The 2013 Clean Air Action (Council, 2013a) identified three megacity clusters as target regions for reducing air pollution: Beijing-Tianjin-Hebei (BTH; $35-41^{\circ}N$, $113.75-118.75^{\circ}E$), Yangtze River Delta (YRD; $29-33^{\circ}N$, $118.75-123^{\circ}E$), and Pearl River Delta (PRD; $21-25^{\circ}N$, $111.25-116.25^{\circ}E$). The more recent plan released in July 2018 (Chinese State Council, 2018b) removed PRD from the list of target regions and added Fenwei Plain (FWP; $33-35^{\circ}N$, $106.25-111.25^{\circ}E$ & $35-37^{\circ}N$, $108.75-113.75^{\circ}E$). Previous studies (X. Zhang et al., 2012) also identified Sichuan Basin (SCB; $27-33^{\circ}N$, $103.75-108.75^{\circ}E$) as one of the major haze regions in China. We present analyses for these five target regions by averaging the data from all sites with more than 70% data coverage for each of the 6 years from of 2013- to 2018. The only continuous record for 2013-2018 in the FWP region is for Xi'an (13 sites). Additional FWP sites outside Xi'an started operating in early 2015 and are consistent with the Xi'an data, as will be shown below.

2.2. Multiple linear regression model

We construct a stepwise multiple linear regression (MLR) model to quantify the effect of meteorology on $PM_{2.5}$ variability. The model fits the deseasonalized and detrended 10-day $PM_{2.5}$ mean time series on the $2^{\circ} \times 2.5^{\circ}$ grid to the five deseasonalized and detrended 10-day mean meteorological variables (WDS, PRECIP, RH, TEM, ~~and~~ V850). The deseasonalized and detrended time series are obtained by subtracting removing the 50-day moving averages from the 10-day mean time series (Tai

et al., 2010). This focuses on synoptic scales of variability and avoids aliasing from common seasonal variations and long-term trends between variables (Shen et al., 2017).

Separate fits of $PM_{2.5}$ to the meteorological variables are done for each $2^\circ \times 2.5^\circ$ grid square and season (DJF, MAM, JJA, SON). The fit has the form:

$$Y_{d,i,s}(t) = \sum_{k=1}^5 \beta_{k,i,s} X_{d,k,i,s}(t) + b_{i,s} \quad (1)$$

where $Y_{d,i}(t)$ is the deseasonalized and detrended $PM_{2.5}$ time series for grid square and season i , and $X_{d,i,k}(t)$ is the corresponding time series for the deseasonalized and detrended meteorological variable $k \in [1,5]$. We fit the regression coefficients $\beta_{i,k}$ and the intercept b_i . The regression is done stepwise, **adding and deleting terms** based on their independent statistical significance to obtain the best model fit (Drapper and Smith, 1998). The fits and the selected meteorological variables differ by location and season but with regional consistency (Table S1). For meteorological variables not in the final MLR model, the regression coefficients $\beta_{i,k}$ in equation (1) are zero.

2.3. Application to 2013-2018 $PM_{2.5}$ trends

We use the MLR model to remove the effect of meteorological variability from the 2013-2018 $PM_{2.5}$ trends, including not only the 10-day synoptic-scale variability but also any interannual variability and 6-year trends. This makes the standard assumption that the same factors that drive synoptic-scale variability also drive interannual variability (Jacob and Winner, 2009; Tai et al., 2012). We thus apply equation (1) to the meteorological anomalies $X_{a,i,k}$, obtained by removing the 6-year means of the 50-day moving averages from the 10-day mean time series. The anomalies **calculated in this manner** are deseasonalized but not detrended. This yields the meteorology-driven $PM_{2.5}$ anomalies $Y_{m,i}$:

$$Y_{m,i}(t) = \sum_{k=1}^5 \beta_{i,k} X_{a,i,k}(t) + b_i \quad (2)$$

Consider now the $PM_{2.5}$ anomaly $Y_{a,i}$ for grid square and season i obtained by deseasonalizing but not detrending the $PM_{2.5}$ data (**by removing the 6-year means of the 50-day moving averages**), in the same way as for the meteorological variables. The residual anomaly $Y_{r,i}$ after removing meteorological influence from the MLR model is given by

$$Y_{r,i}(t) = Y_{a,i}(t) - Y_{m,i}(t) \quad (3)$$

The residual is the component of the anomaly that cannot be explained by the MLR meteorological model **and we will refer to it as the meteorology-corrected data**. It includes noise due to limitations of the MLR model and other factors, but also a long-

term trend over the 6-year period that we can attribute to changes in anthropogenic emissions. The same approach was recently applied by Li et al. (2019) to separate anthropogenic and meteorological drivers of ozone trends in China.

3. Results and Discussion

3.1. PM_{2.5} trends in China, 2013-2018

Figure 1 shows annual mean observed PM_{2.5} concentrations from the CNEMC over China for 2013 and 2018, and the linear regression trends on the 2°×2.5° grid based on the PM_{2.5} anomalies $Y_{a,i}(t)$ including effects of both changing emissions and meteorology. In 2013, PM_{2.5} across most of China was well above the Chinese national air quality standard (annual mean of 35 μg m⁻³). BTH and FWP (Xi'an) had the highest PM_{2.5} among the five target regions, with annual average concentrations of 108 ± 34 μg m⁻³ (standard deviation describes variability of the annual average across sites in the region) and 108 ± 11 μg m⁻³ respectively, followed by SCB (71 ± 17 μg m⁻³), YRD (67 ± 12 μg m⁻³), and PRD (47 ± 7 μg m⁻³). PM_{2.5} decreased dramatically from 2013 to 2018, by 34-49% for the five target regions. Mean 2018 concentrations were 55 ± 13 μg m⁻³ in BTH, 62 ± 4 μg m⁻³ in FWP (Xi'an), 40 ± 6 μg m⁻³ in SCB, 40 ± 7 μg m⁻³ in YRD, and 31 ± 5 μg m⁻³ in PRD.

Figure 2 shows the 2013-2018 relative trends of annual mean PM_{2.5} for the five target regions, along with the corresponding trends of SO₂, NO₂, and CO concentrations measured at the same sites. Also shown in the bottom panels are the MEIC inventory trends in emissions of primary PM_{2.5}, SO₂, NO_x, NH₃, and CO for 2013-2017. The PM_{2.5} observations show steady decreases for BTH, YRD, and SCB. PRD flattens out in 2015-2017 before decreasing again in 2018. FWP (Xi'an) decreases sharply by 47% from 2013 to 2015 but rebounds in 2015-2017 before decreasing again in 2018. Trends at other FWP sites that became operational in early 2015 are similar to Xi'an. We argue in Section 3.3 that the 2015-2017 flattening at PRD and the anomalously 2013-2015 sharp decrease and 2015-2017 rebound at FWP are driven by meteorology.

We see from Figure 2 that only SO₂ has a decrease steeper than PM_{2.5}, indicating that SO₂ emission controls have been a major driver of the PM_{2.5} trend (Lang et al., 2017; Shao et al., 2018). The overall SO₂ decrease for the five regions is 57-76% from 2013 to 2018. The SO₂ decrease is quantitatively consistent with the decrease of SO₂ emissions estimated by MEIC (Zheng et al., 2018). This drastic cut of China SO₂ emissions is due to installation of scrubbers at coal-fired power plants (Siwen et al., 2015; Karplus et al., 2018; Silver et al., 2018), elimination of small coal boilers, improvement of coal quality (Zheng et al., 2018), and switch from residential coal to cleaner fuels (Zhao et al., 2018). We also see a significant decrease in CO of 18-43% for the five regions from 2013 to 2018, again consistent with the MEIC emission inventory and suggesting a reduction in organic PM_{2.5} emissions. Primary PM_{2.5} emissions in the MEIC inventory decreased at a rate comparable or

steeper than CO. [Trends in China PM_{2.5}, SO₂, and NO₂ presented here are consistent with previous studies \(Silver et al., 2018; Ma et al., 2019\) that cover a shorter time period than 2013-2018.](#)

Figure 3 shows the time series of monthly mean PM_{2.5} for the five target regions, illustrating the seasonal and interannual variability. All regions show winter maxima that can be mostly attributed to meteorology including shallower mixing depth, lower precipitation, and increased stagnation in winter (X. Wang et al., 2018). Residential heating emissions in winter also contribute to the seasonality in northern China (J. Liu et al., 2016b; Xiao et al., 2015). There is a large interannual variability, particularly in winter, that must be largely driven by meteorology. Studies for BTH have shown that high PM_{2.5} in winter months is associated with weak southerly winds, low mixing depths, and high relative humidity (R. Zhang et al., 2014; Chang et al., 2016; K. Li et al., 2018; Shao et al., 2018). The relatively clean 2017-2018 winter was due in part to a higher frequency of northerly flow and associated ventilation (Administration, 2018; Yi et al., 2019). In addition, particularly aggressive actions by the government to restrict coal use that winter may have played a role in reducing PM_{2.5} levels (X. Zhang et al., 2019).

3.2. Meteorological influence on PM_{2.5}

Figure 4 shows the correlations of 10-day PM_{2.5} concentrations with the individual meteorological variables used in the MLR model. Wind speed is negatively correlated with PM_{2.5}, as would be expected from ventilation, except in areas of the north where wind promotes dust formation (Lyu et al., 2017; X. Wang et al., 2004). Precipitation is also generally negatively correlated with PM_{2.5}, as one would expect from scavenging (Chen et al., 2018). The positive correlation between precipitation and PM_{2.5} over north China in spring is likely a result of high RH associated with precipitation in adjacent days.

Correlation between RH and PM_{2.5} is positive over northern China, especially in winter, and negative over southern China, especially in summer. The positive correlation between PM_{2.5} and RH over northern China in winter has been reported by previous studies and attributed in part to the role of aqueous-phase aerosol chemistry in driving secondary PM_{2.5} formation (Zheng et al., 2015; He et al., 2018; Song et al., 2019; Pendergrass et al., 2019; Tie et al., 2017). The negative correlation of PM_{2.5} with RH over south China likely reflects the association of high RH with precipitation and onshore wind, which facilitate PM_{2.5} wet removal and ventilation (Zhu et al., 2012; Leung et al., 2018).

Temperature has a positive correlation with PM_{2.5} year round over most of China (Y. Wang et al., 2014; Leung et al., 2018), even though there is no strong direct dependence of PM_{2.5} on temperature (Jacob and Winner, 2009). The correlation likely reflects the covariation of temperature with the other meteorological variables (Tai et al., 2012; Zhu et al., 2012). A possible explanation for the negative correlation with temperature in summer over North China Plain could be the volatilization of

ammonium nitrate at high temperature (Kleeman, 2008). V850 shows strong positive correlations with winter $PM_{2.5}$ over most of China, and strong negative correlations with summer $PM_{2.5}$ over south China, especially for the Pearl River Delta.

Figure 5 (left panel) describes the ability of the MLR model to account for $PM_{2.5}$ variability in relation to wind speed, precipitation, RH, temperature, and V850 as potential predictor variables. Results are presented as the coefficients of determination R^2 (fraction of variance explained) between observed and model $PM_{2.5}$ in the detrended deseasonalized time series. The R^2 values have been adjusted to account for different numbers of significant explanatory terms (predictor variables). R^2 values for the five target regions are: 0.59 (BTH), 0.46 (YRD), 0.65 (PRD), 0.65 (SCD), and 0.41 (FWP). The right panel of Figure 5 shows the meteorology-corrected $PM_{2.5}$ trends after removal of meteorological variability predicted by the MLR model, i.e., the trends in the residuals $Y_{r,i}(t)$ in equation (3). The meteorology-corrected decreasing trend averaged across China is $-4.6 \mu\text{g m}^{-3} \text{a}^{-1}$, 12% weaker than the trend in the original data (Figure 1). We elaborate below for the five target regions.

3.3. Meteorology-corrected $PM_{2.5}$ trends for the five target regions

Figure 6 shows the 10-day mean $PM_{2.5}$ anomalies in the deseasonalized (but not detrended) data for the five target regions ($Y_a(t)$ in Section 2.3). Also shown is the meteorological component $Y_m(t)$ derived from the MLR meteorological model, and the residual $Y_r(t)$ (meteorology-corrected, equation (3)) whose long-term trend can be interpreted as due to changes in anthropogenic emissions. The $PM_{2.5}$ anomalies show large features on ten-day basis that can be mostly captured by the MLR model. The residual meteorology-corrected time series is much smoother, as depicted by the narrower 95% confidence intervals in the anthropogenic residual trends than in the original observed trends. The meteorology-corrected trends differ by 3% (YRD) to 27% (SCB) from the observed trends. The YRD trend reflects a significant contribution from the December 2013 outlier, which reflects unfavorable meteorological conditions (Figure S2) that are not adequately captured by the MLR model. If we exclude this outlier month from the time series, the observed YRD trend becomes $-5.7 \pm 0.9 \mu\text{g m}^{-3} \text{a}^{-1}$ and the meteorology-corrected trend becomes $-5.9 \pm 0.7 \mu\text{g m}^{-3} \text{a}^{-1}$.

Most remarkably, it appears that the 2015-2017 flattening in the PRD and 2015-2017 increase in the FWP can be mostly attributed to meteorological variability as resolved by the MLR model, rather than to emissions. The trend in the residual is more consistent with a steady 2013-2018 anthropogenic decrease in both regions. The MLR model shows that meteorology accelerated the $PM_{2.5}$ decline in PRD and FWP from 2013 to 2015, and contributed partly to the 2015-2017 $PM_{2.5}$ rebound over FWP. In particular, the high $PM_{2.5}$ anomalies in PRD in 2013 and early 2014 are driven by anomalously low V850, and the low $PM_{2.5}$ in winter 2015-2016 is associated with anomalously high southerly flow and precipitation (Figure S4). The

low PM_{2.5} in FWP in the winter 2014-2015 is associated with anomalously high wind speed, low RH, and low temperature, while the high anomalies in the winter 2016-2017 are associated with anomalously low wind speed, high RH, and high temperature (Figure S5).

4. Conclusions

Observations of fine particulate matter (PM_{2.5}) pollution in China from the extensive CNEMC network established in 2013 show large 2013-2018 decreases ~~in apparent response to driven by~~ emission controls with complicating influences from meteorology. Here we used a stepwise multiple linear regression (MLR) meteorological model to investigate and quantify the meteorological contributions to these 6-year trends.

The CNEMC observations show 34-49% decreases in PM_{2.5} in the five megacity clusters targeted by the Chinese government's Clean Air Action to reduce anthropogenic emissions. Concurrent observations of SO₂, CO, and NO₂ are qualitatively consistent with these PM_{2.5} decreases being driven by drastic cuts in emissions from coal combustion. At the same time, there is large interannual variability driven by meteorology particularly in winter when PM_{2.5} is highest.

We used the stepwise MLR meteorological model to relate PM_{2.5} anomalies across China to wind speed, precipitation, relative humidity (RH), temperature, and meridional velocity at 850 hPa (V850) as potential predictors. The model accounts for ~50% of the variance in the deseasonalized detrended PM_{2.5} data, including 41-65% for the five megacity clusters. Application to the PM_{2.5} time series shows that meteorological variability contributed significantly to the 6-year trends across China and in the megacity clusters. Removing meteorological variability as given by the MLR model also reduces the uncertainty in the trend that can be attributed to emission controls. We refer to the data series after removal of meteorological variability as the meteorology-corrected data. Thus the 2013-2018 PM_{2.5} decrease for Beijing-Tianjin-Hebei is $-9.3 \pm 1.8 \mu\text{g m}^{-3} \text{ a}^{-1}$ in the original data and is 14% weaker in the ~~meteorology-corrected data~~ ($-8.0 \pm 1.1 \mu\text{g m}^{-3} \text{ a}^{-1}$) ~~in the meteorology-corrected data~~. For the Sichuan Basin where the meteorological correction is particularly large, the PM_{2.5} decrease is $-6.7 \pm 1.3 \mu\text{g m}^{-3} \text{ a}^{-1}$ in the original data and is reduced by 27% to $-4.9 \pm 0.9 \mu\text{g m}^{-3} \text{ a}^{-1}$ in the meteorology-corrected data ~~$4.9 \pm 0.9 \mu\text{g m}^{-3} \text{ a}^{-1}$ in the meteorologically corrected data~~. The average 2013-2018 PM_{2.5} decrease over our study domain is $-5.2 \mu\text{g m}^{-3} \text{ a}^{-1}$ in the original data (Figure 1 (right panel)); and is reduced by 12% to $-4.6 \mu\text{g m}^{-3} \text{ a}^{-1}$ in the meteorology-corrected data (Figure 5 (right panel)).

Observations for the 2015-2017 period indicate a flattening of the PM_{2.5} trend in the Pearl River Delta and an increase in the Fenwei Plain. We find from the MLR model that these 3-year trends can be explained by meteorological variability (including particularly steep 2013-2015 decreases) rather than by relaxation of emission controls.

Data availability. All of the measurements, reanalysis data are openly available for download ~~form~~ [from](#) the websites given in the main text. The anthropogenic emission inventory is available from www.meicmodel.org, and for more information, please contact Qiang Zhang (qiangzhang@tsinghua.edu.cn).

Competing interests. The authors declare that they have no conflict of interest.

255 **Author contributions.** SXZ, DJJ, and HL designed the study. SXZ developed the model, performed the simulations and analyses. XW, LS, KL, YZZ, and TLZ helped with scientific interpretation and discussion. KG helped with pollutants data processing. SXZ and DJJ wrote the manuscript and all authors provided input on the paper for revision before submission.

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Annual mean PM_{2.5} concentrations and trends, 2013-2018

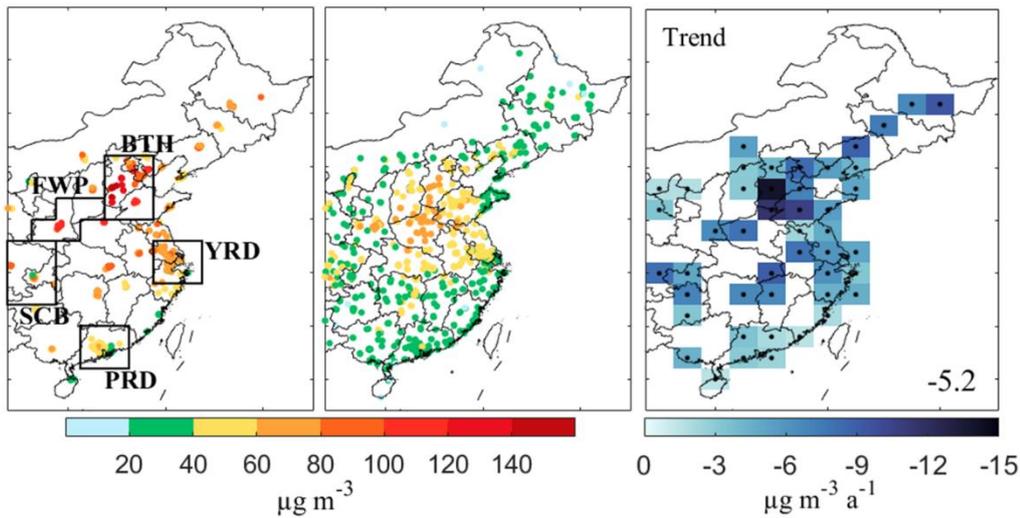


Figure 1 Annual mean PM_{2.5} concentrations in China from the CNEMC network. Left and middle panels show values for 2013 and 2018 for sites with more than 70% data coverage for ~~each~~ the corresponding year. The right panel shows the ordinary linear regression trends on a 2°x2.5° grid for sites with more than 70% data coverage for each of the ~~five~~ six years from 2013 to 2018. The trends are based on the timeseries of 10-day mean anomalies as described in the text. Polygons in the left panel define the four target regions of the Clean Air Action (Beijing-Tianjin-Hebei (BTH; 35–41°N, 113.75–118.75°E), Yangtze River Delta (YRD; 29–33°N, 118.75–123°E), Pearl River Delta (PRD; 21–25°N, 111.25–116.25°E), and Fenwei Plain (FWP; 33–35°N, 106.25–111.25°E & 35–37°N, 108.75–113.75°E)), to which we add Sichuan Basin (SCB; 27–33°N, 103.75–108.75°E). Number inset in the right panel is the trend in mean PM_{2.5} over the study region (21–41°N, 103.75–123°E). Dots in the right panel indicate grid squares with significant trends ($p < 0.05$).

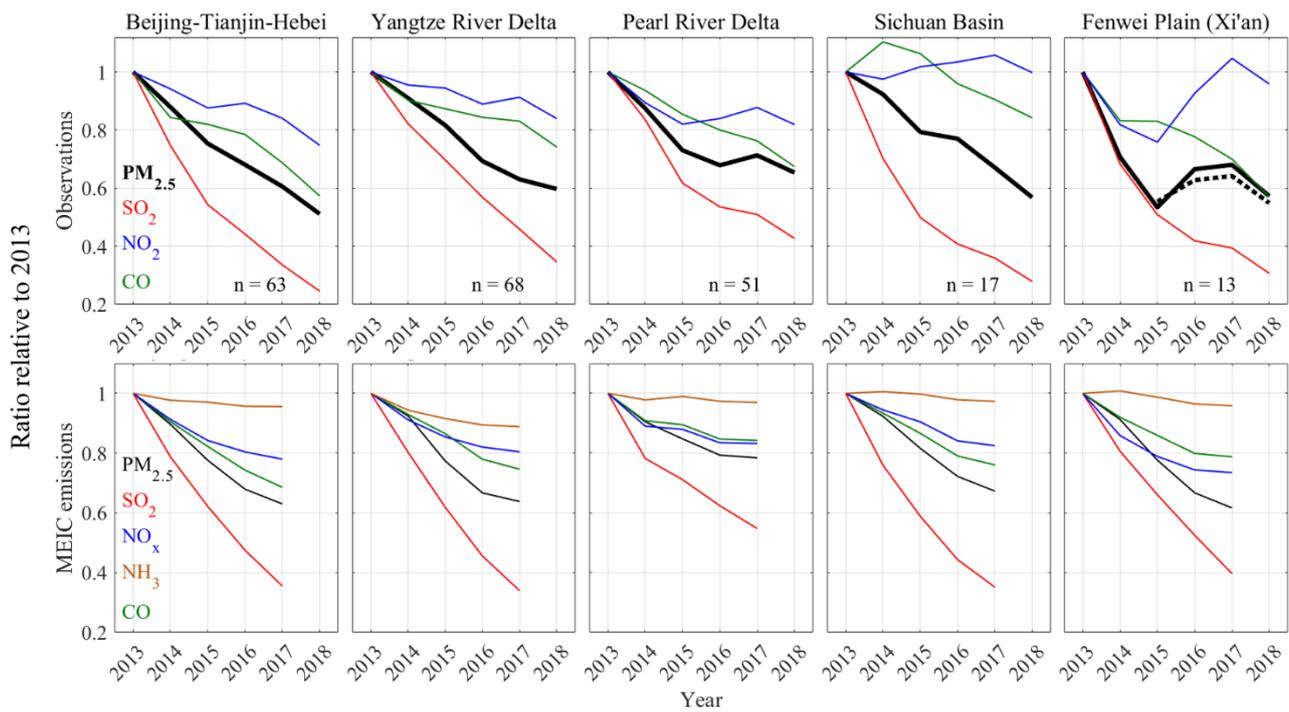
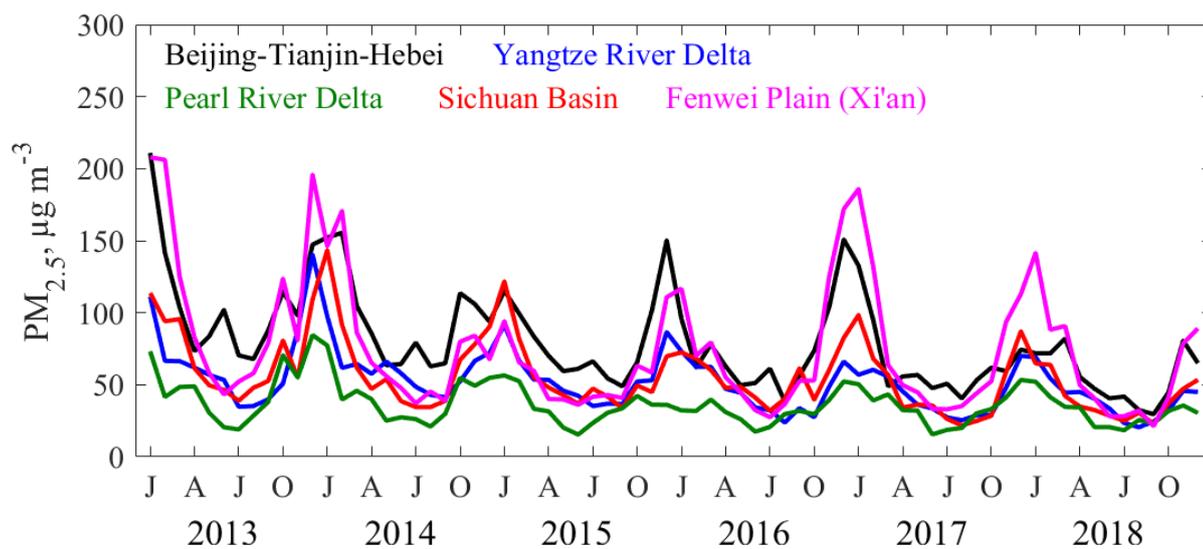


Figure 2. Relative trends of 2013-2018 observed concentrations and 2013-2017 MEIC emission estimates for the five target regions of Figure 1. Values are annual means referenced to 2013. The observed concentrations are averaged over all sites in each region with at least 70% data coverage for each year. The number of sites for each region is indicated. Fenwei Plain trends are for Xi'an as other sites did not start operating until early 2015. Post-2015 relative PM_{2.5} trends at these other sites are shown as the dashed line.

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470 **Figure 3** 2013-2018 time series of monthly mean $PM_{2.5}$ concentrations over the five target regions. Values are averages from all sites in the region with over 70% data coverage for each year of the six years.

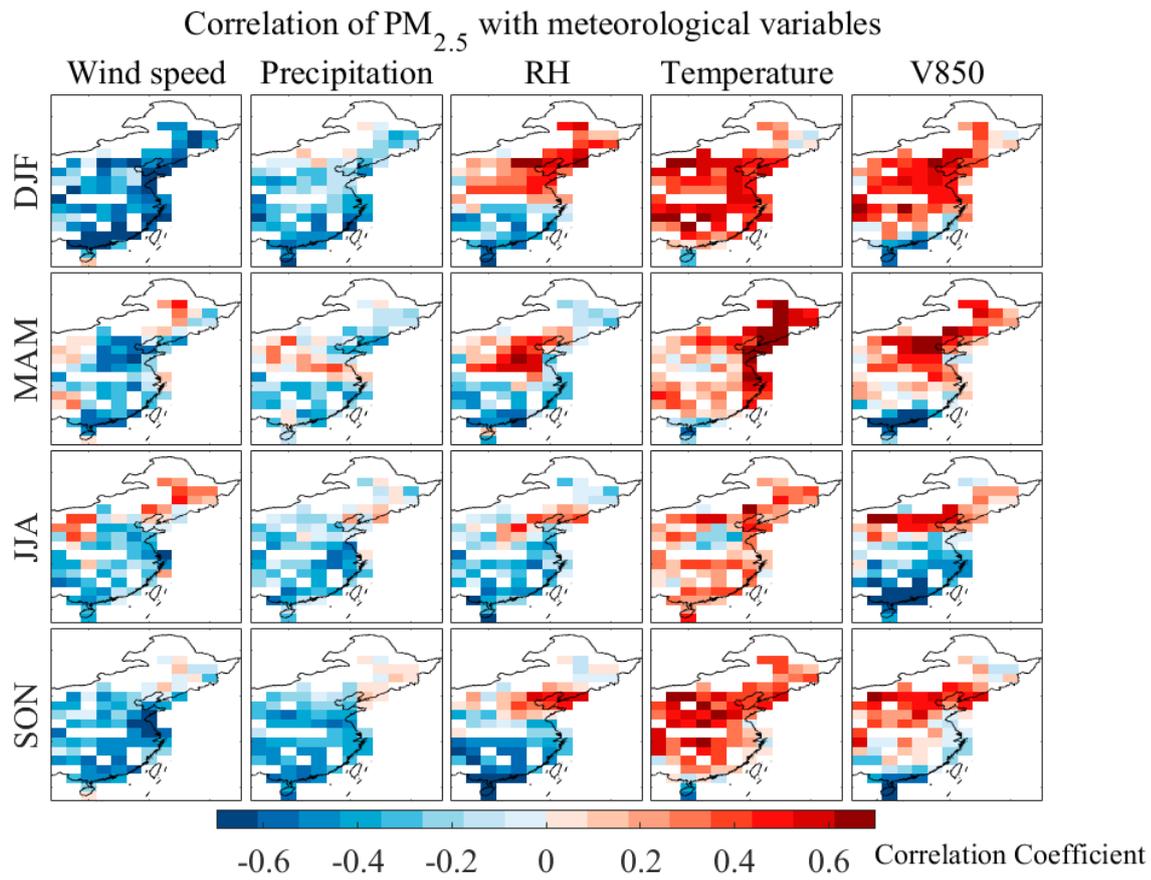
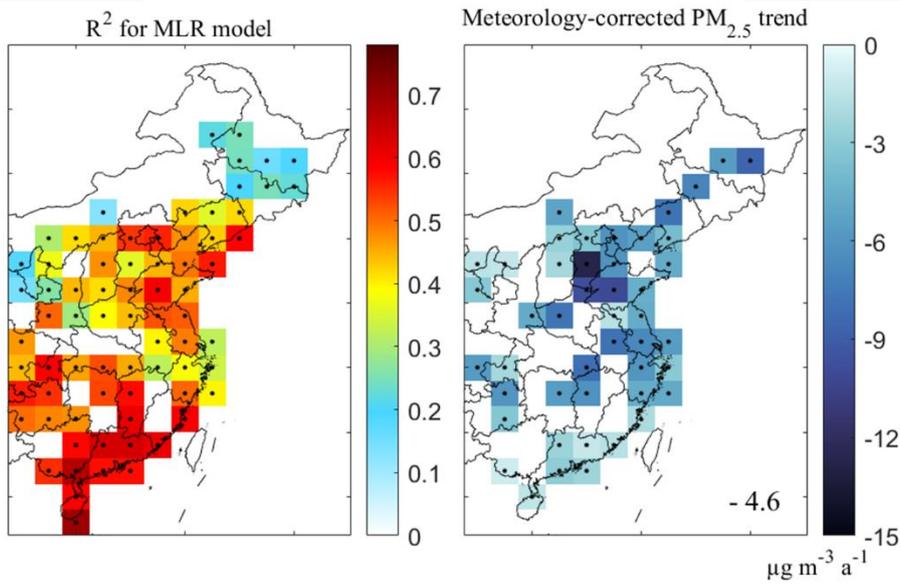


Figure 4 Correlation coefficients (r) of PM_{2.5} concentration with the individual meteorological variables used in the MLR model: surface wind speed (m s⁻¹), precipitation (mm d⁻¹), relative humidity (RH; %), surface air temperature (°C), and 850hPa meridional wind velocity (m s⁻¹) for different seasons in China. The correlations are based on 10-day average observations on a 2°×2.5° grid.

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Meteorological influences on 2013-2018 PM_{2.5} trends



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Figure 5 Resolving meteorological influences on PM_{2.5} 2013-2018 trends in China. The left panel shows the fraction of detrended and deseasonalized variance in 10-day PM_{2.5} means explained by the stepwise multi linear regression (MLR) meteorological model. The right panel shows the meteorology-corrected trends, to be compared to the trends in the original data shown in Figure 1. Number inset in the right panel is the trend in mean PM_{2.5} over the study region (same definition as in Figure 1). Dots indicate significant correlations ($p < 0.05$) in the left panel and significant trends ($p < 0.05$) in the right panel.

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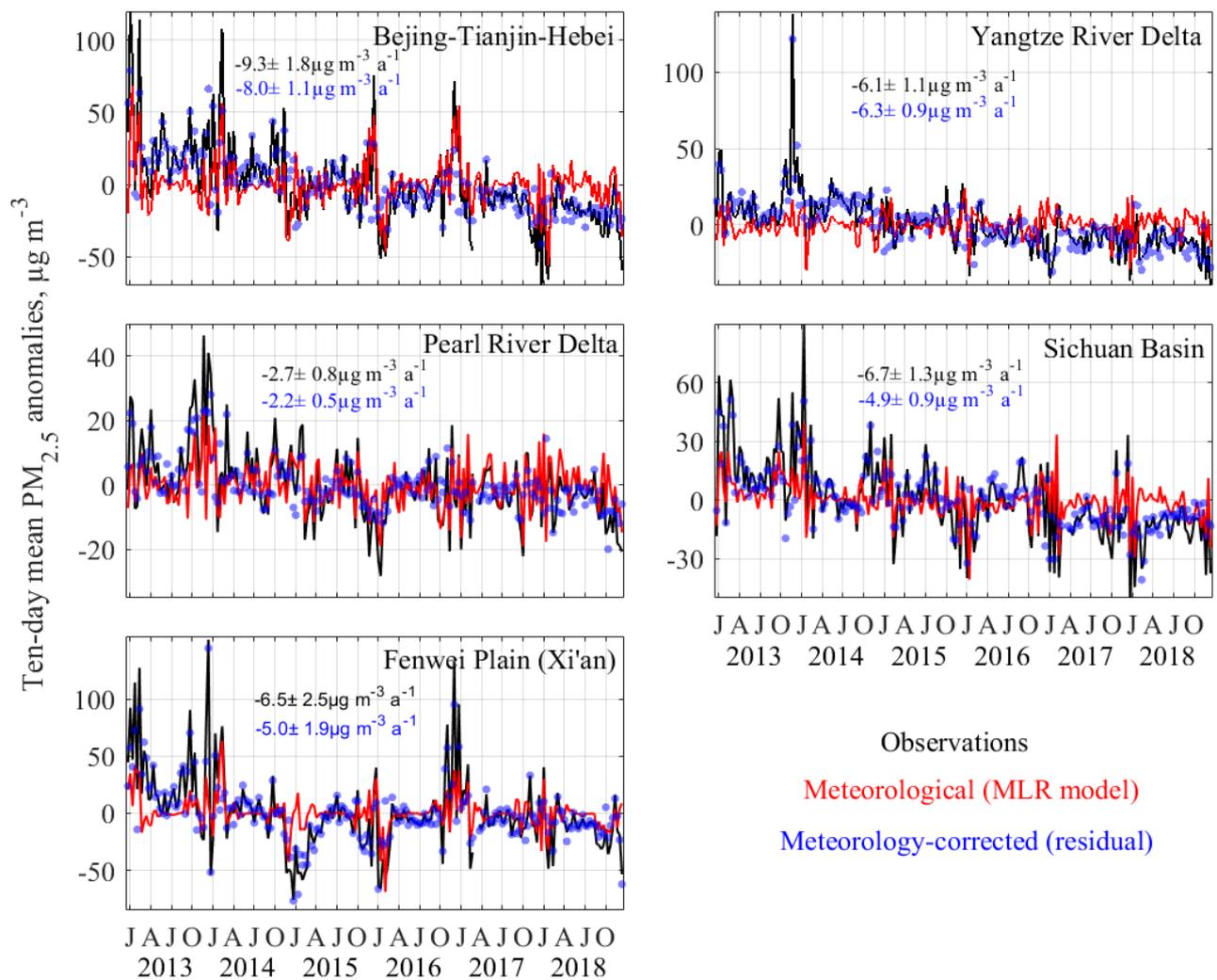


Figure 6. Time series of 2013-2018 PM_{2.5} 10-day mean anomalies for the five target regions of Figure 1. The anomalies are relative to the 2013-2018 means. The data are averaged over all measurement sites in each region with at least 70% of data coverage for each year (same as for Figure 2). The meteorological contribution to the anomalies as diagnosed from the MLR model is shown in red. The long-term trend in the meteorology-corrected residual in blue (equation (3)) is interpreted as driven by changes in anthropogenic emissions. Values inset each panel are the ordinary linear regression trends with 95% confidence intervals obtained by the bootstrap method.

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