Response to Referee Comments on “Interpreting Space-Based Trends in Carbon Monoxide”

Reviewer comments are in blue, and the response in black.

Response to Reviewer 1:

The manuscript describes a small series of numerical experiments to understand previously reported trends in column CO from the MOPITT satellite instrument. It is a thoughtful interpretation of the MOPITT data. The manuscript would benefit from more detail in places, detailed below, but it is suitable for publication in ACP.

We thank the referee for the review and respond to individual comments below.

Detailed comments:

1) The manuscript title would benefit from being more specific.

We updated the title to: “Interpreting Space-Based Trends in Carbon Monoxide with Multiple Models”

2) MACCity or MACCITY? Be consistent.

We now use MACCity throughout the manuscript.

3) This reader thought the abstract would benefit from being punchier. What is the punchline? Is it that getting column ozone right is important for understanding column CO trends?

We added the following sentences to the abstract to emphasize our conclusions: “This demonstrates that biases in a model’s average concentrations can influence the interpretation of the temporal trend compared to satellite observations.” and “These results demonstrate that accurately simulating variability in the ozone column is important for simulating and interpreting trends in CO.”

4) Line 71. It would be useful for the reader to be more specific about assumptions/data used to build MACCCity. Maybe a few sentences so the reader is not required to immediately chase up details elsewhere.

We added the following details in section 2.2: “The MACCity inventory linearly interpolates the decadal anthropogenic emissions from the ACCMIP inventory (Lamarque et al., 2010) for 2000, and the RCP8.5 emissions for 2005 and 2010, to each year in between. The MACCity biomass burning emissions have year-to-year variability based on the GFED-v2 (van der Werf et al., 2006) inventory.”

5) Line section 2.1. This reader believes it would be useful to show an example MOPITT averaging kernel. Just for my own curiosity, is there a difference in the
averaging kernel over eastern China during 2010?

We added a supplemental figure showing the column averaging kernels for 2010 and two other years. We added this discussion to Section 2.1: “Supplemental Figure S1 shows the MOPITT column averaging kernels averaged over four regions. The column averaging kernels depend on the observed scene, and vary year to year as well as seasonally. The dependence of the column averaging kernels on the CO mixing ratio profile (Deeter, 2009) explains the high values in the lower troposphere over eastern China in winter.” Although the kernel over eastern China varies year to year, we did not find 2010 to be fundamentally different from other years.

6) Line 90: What did the authors assume when calculating the autocorrelation?

We clarify that we calculated the autocorrelation for a 1-month lag.

7) Line 91: By deseasonalizing the column data by fitting sines/cosines the authors are implicitly assuming stationarity of these data. Do the authors believe this is a valid assumption over a decade-long time series during which the phase and amplitude of the seasonal cycle changes?

Although there are anomalies in particular months that cause year-to-year variability in the seasonal cycle, we did not find a systematic change in the phase and amplitude of the seasonal cycle.

8) Line 93: Months with insufficient data? How do the authors define this criterion?

We now clarify: “Months with no MOPITT data or only a few days of MOPITT data are excluded from the trend analysis. This includes May-August of 2001 and August-September of 2009.”

9) Line 95: Do the authors sample the model along the orbit tracks?

We use the level 3 data, which is a gridded product. We added this information in section 2.1.

10) Minor comment: a few instances where the references should be inline but are not. I expect the typesetting process will pick this up.

Fixed.

11) Section 2.2: Would be useful if the authors provide some regional emission estimates of CO, particularly for pertinent regions.

We added the following information: “From 2000 to 2010, CO emissions in the MACCity inventory decreased from 31 to 11 Tg yr⁻¹ over the eastern U.S., from 97 to 59 Tg yr⁻¹ over Europe, and increased from 56 Tg to 72 Tg yr⁻¹ over eastern China.”
12) Equation 1. The usual convention is lower-case bold typeface for vectors and upper-case bold typeface for matrices.

We modified the equation to reflect this convention.

13) This reader is confused why the authors included two sets of statistics: 2000-2010 and 2000-2011. If the results from these two sets had been significantly different I would have probably suggested a major re-analysis of the data. But they are not so I suggest (and only suggest) the authors summarize the value of the additional data in a few sentences and report only 2000-2010.

We updated the text and figures to report only values for 2000-2010, and removed the supplemental figure, which showed results for 2000-2011. We added the following sentence to Section 2.2: “Some simulations were available through 2011, while others ended in 2010. We therefore report results for 2000-2010, but note that extending the analysis through 2011 does not alter the conclusions.”

14) Line 186 and elsewhere: The authors won’t need reminding that a model-data correlation r of just over 0.7 is required for the model to describe 50% of the observed variation. In some places the authors extol correlations of X (much less than 0.7) while in other places the author extol the squared correlation values of Y.

We agree that many of the r values shown in Table 2b are below 0.7 and thus imply that the simulations are capturing less than half of the variance. However, we find it useful to report and discuss these r values to indicate the relative performance of different simulations. We updated this table to indicate which correlations are statistically significant. We updated the text in Section 3.3 to report r values rather than r² for consistency with the rest of the paper.

15) Line 265: What is responsible for the observed and model total column anomaly in 2010?

Steinbrecht et al. (2011) attribute the 2010 anomaly in northern midlatitude ozone observations to a combination of an unusually strong negative Arctic Oscillation and North Atlantic Oscillation and the easterly phase of the quasi-biennial oscillation. We added this information in Section 3.3.

16) Line 271: “can be” or “is”?

We changed this to “is partially”.

Response to Referee 2:

The authors have used a series of chemistry-climate and chemical transport model simulations to understand the negative trends in CO observed by MOPITT. They find that the negative trend in the bottom-up inventories reproduce the trend observed
over North America and Europe, but is incapable of capturing the negative trend observed over China. They attributed the discrepancies between the modeled and observed trend in CO over China to changes in the MOPITT vertical sensitivity and to biases in the modeled ozone abundances, which produces a bias in modeled OH and, thus, CO. The paper is well written and the authors did a careful analysis of the trends in the models. I would recommend publication in ACP after the authors have revised the manuscript to address my comments below.

We appreciate the thoughtful review and respond to comments below.

Comments

1) Line 85: The Level 3 MOPITT data are daily or monthly gridded data. Are the authors using the daily gridded data here? Are they using nighttime and daytime, or just daytime data? If they are using Level 3 data, how do they compare the model to the MOPITT data. The model should be sampled at the MOPITT observation locations and times when transforming the model with the MOPITT averaging kernels and a priori profiles. They need to better explain in the paper how this is done.

We now clarify in section 2.1 that we are using the monthly gridded daytime data, and that the level 3 product includes the averaging kernel and a priori for each grid box. In section 2.2, we added that we regrid the model results to the MOPITT grid. We expect the error from using monthly mean simulated CO instead of sampling at the overpass time to be small since CO does not have a large diurnal cycle. Our analysis includes a free-running CCM simulation as well as CTM simulations. The meteorology of the free-running CCM will not match up with the observed for individual days, and our focus is on monthly and interannual rather than daily variability, so we chose to use the monthly mean MOPITT product. Martinez-Alonso et al [2014] demonstrated that there is only a small bias from using gridded average averaging kernels rather than the kernels of individual retrievals.

2) Lines 137 – 141: In giving Equation (1), the authors should explain that the MOPITT retrievals are with respect to the log of the mixing ratio and they should also cite Deeter (2009) “MOPITT (Measurements of Pollution in the Troposphere) Validated Version 4 Product User’s Guide” for providing guidance for calculating the column averaging kernels from the averaging kernels that is with respect to the log of the mixing ratio.

We added the following line to Section 2.2: “The column averaging kernel is calculated from the standard averaging kernel matrix, which is based on the log of the CO concentration profile, following the method of Deeter (2009):

\[ a_j = \frac{K}{\log_{10}e} \sum \Delta p_i v_{rtv,i} A_{ij} \] (2)

where \( \Delta p_i \) and \( v_{rtv,i} \) are the pressure thickness and retrieved CO concentration, respectively, of level i, A is the standard averaging kernel matrix, and \( K = 2.12 \times 10^{13} \)
molec cm$^{-2}$ hPa$^{-1}$ ppb$^{-1}$.

3) Lines 201-203: The authors state here that the discrepancy is “driven largely by the failure of the simulations to capture the 2008 dip,” but the models are also strongly biased in 2010, for example (Fig. 2f). Indeed, this 2010 bias is the focus of the ozone analysis in Figs. 3 and 4.

We added: “leading to an overestimate that continues through 2010”.

4) Please state the regional boundaries for the regions considered in Fig. 2.

We added this information to the caption of Fig. 2.

5) Lines 226 – 244: I find this discussion confusing. I understand how time-dependent variations in the vertical sensitivity of the MOPITT retrievals could contribute to trends in the data. But I don’t understand how, as stated on Lines 228-321, the bias in the modeled CO can produce an artificial trend. It seems to me that there are two possible way this could happen:

a) Are the authors suggesting that changes in the vertical distribution of the model bias, combined with the varying vertical sensitivity of the MOPITT retrievals, produces an artificial trend when the model is convolved with the averaging kernels and a priori profile?

b) If the model bias is constant in time, then the convolved modeled columns should exhibit the same trend as the data. The presence of a fixed bias in the model together with temporally varying averaging kernels should only impact the trend if the biased model state is so far from the a priori MOPITT state that the linearization assumption in Equation (1) is invalid i.e., the averaging kernels do not accurately capture the sensitivity of the retrieval between the modeled state and the MOPITT a priori. Is this what the authors are trying to say on lines 226-228?

The authors need to explain more clearly how the bias in the model could be contributing to an artificial trend.

As the reviewer suggests in a), the bias in CO varies with altitude, so if the vertical sensitivity described by the averaging kernel changes to e.g. place more weight on higher altitudes, this will change the value of the convolved CO column even if there were no changes in the CO profile. But even if the bias were constant with altitude, changes in the averaging kernel result in more or less weight placed on the a priori versus the CO simulated by the model. Thus, a difference between the a priori and the model means that placing more (or less) weight on the a priori will change the resulting value of $C_{\text{sim}}$. Since the a priori profiles and columns are constant in time, taking the time derivative of equation 1 yields:

$$\frac{\partial C_{\text{sim}}}{\partial t} = a \left( \frac{\partial x_{\text{mod}}}{\partial t} \right) + \frac{\partial a}{\partial t} (x_{\text{mod}} - x_0)$$
The second term on the right hand side shows that the larger the bias between the modeled CO and the a priori, the larger the impact of the changing averaging kernel. We now discuss this in Section 3.2.

6) Lines 261-262: Yes, we expect that anomalies in OH to be inversely related to anomalies in total ozone, however, the OH and ozone anomalies do not seen to be strongly correlated in Fig. 3. It would be helpful if the authors gave the correlation coefficient between the two quantities for different latitude bands in the tropics and extratropical northern hemisphere.

We added the following sentence to this discussion: “The correlation coefficient between OH and column ozone is -0.53 for the 15°S-15°N average, -0.72 for the 15°-25°N average, and -0.75 for the 30°-60°N average.”

7) Lines 263-264: The ozone column anomaly in Fig. 3 is in the extratropical northern hemisphere, mainly in early (Jan-Mar) 2010. Although the global, annual mean CO lifetime is 1-2 months, in the extratropics in winter it could be longer than a season. If that is the case, it is unclear to me how the changes in OH in early 2010 could drive such large changes in CO between 30-60N in winter.

The early 2010 OH anomaly occurs in the northern tropics as well as the extratropics. We added the following sentence to highlight this: “This OH anomaly extends from the northern tropics to the midlatitudes.” Northern midlatitude (30-60N) CO and OH are anticorrelated with an r value of -0.69. This r value increases to -.78 if we apply a three-month smoothing, reflecting the longer lifetime of CO. However, since the lifetime of CO is several months, we do not expect a one-to-one correspondence between CO anomalies and OH anomalies, and we added a sentence stating this. We also clarified figure 3 by adding the units to each panel.

8) Figure 4: The 2010 ozone anomaly is about 5%. What are the altitude ranges that are contributing to this bias in the column? Are these changes mainly in the UTLS?

The 2010 anomaly in both the SBUV data and the model is driven by ozone at pressures higher than 25hPa. The vertical resolution of the SBUV data does not allow us to determine the specific altitude of the bias. However, comparison to MLS data shows that GMI has a high bias in lower stratospheric ozone (pressures greater than 50 hPa) in the first half of 2010.

Technical comments

1. Line 64: Change “results of (Li and Liu, 2010)” to “results of Li and Liu (2010)”

Fixed

2. Figure 2: It difficult to see the seven different lines in each panel. If the authors remove the titles on the y-axes on the panels in the right column and reduce the spacing between the panels, it may be possible to enlarge each panel to make the
plots more legible.

We updated this figure as suggested.
Interpreting Space-Based Trends in Carbon Monoxide with Multiple Models


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Abstract
We use a series of chemical transport model and chemistry climate model simulations to investigate the observed negative trends in MOPITT CO over several regions of the world, and to examine the consistency of time-dependent emission inventories with observations. We find that simulations driven by the MACCity inventory, used for the Chemistry Climate Modeling Initiative (CCMI), reproduce the negative trends in the CO column observed by MOPITT for 2000-2010 over the eastern United States and Europe. However, the simulations have positive trends over eastern China, in contrast to the negative trends observed by MOPITT. The model bias in CO, after applying MOPITT averaging kernels, contributes to the model-observation discrepancy in the trend over eastern China. This demonstrates that biases in a model’s average concentrations can influence the interpretation of the temporal trend compared to satellite observations. The total ozone column plays a role in determining the simulated tropospheric CO trends. A large positive anomaly in the simulated total ozone column in 2010 leads to a negative anomaly in OH and hence a positive anomaly in CO, contributing to the positive trend in simulated CO. These results demonstrate that accurately simulating variability in the ozone column is important for simulating and interpreting trends in CO.

1. Introduction
Carbon monoxide (CO) is an air pollutant that contributes to ozone formation and affects the oxidizing capacity of the troposphere (Thompson, 1992; Crutzen, 1973). Its primary loss is through reaction with OH, which leads to a lifetime of 1-2 months (Bey et al., 2001) and makes CO an excellent tracer of long-range transport. Both fossil fuel combustion and biomass burning are major sources of CO. The biomass burning source shows large interannual variability (van der Werf et al., 2010), while fossil fuel emissions typically change more gradually. The time-dependent MACCity inventory (Granier et al., 2011) shows decreases in CO emissions from the United States and Europe from 2000 to 2010 due to increasing pollution controls, but increases in emissions from China. MACCity emissions for years after 2000 are based on the Representative Concentration Pathway (RCP) 8.5 (Riahi et al., 2007). The REAS (Kurokawa et al., 2013) and EDGAR4.2 (EC-JRC/PBL, 2011) inventories also show increasing CO emissions from China. The bottom-up inventory of Zhang et al. (2009) shows an 18% increase in CO emissions from China from 2001 to 2006, and Zhao et al. (2012) estimate a 6% increase between 2005 and 2009. However, there is considerable uncertainty in bottom-up inventories, and comparison of model hindcast simulations driven by bottom-up inventories with observations provides an important test of the time-dependent emission estimates.

Space-based observations of CO are now available for over a decade and show trends at both hemispheric and regional scales. Warner et al. (2013) found significant negative trends in both background CO and recently emitted CO at 500 hPa over southern hemisphere oceans and northern hemisphere land and ocean in Atmospheric Infrared Sounder (AIRS) data. Worden et al. (2013) calculated trends in the CO column from several thermal infrared (TIR) instruments including MOPITT and AIRS. They found statistically significant negative trends over Europe, the eastern United States, and China for 2002-2012. He et al. (2013) also report a negative trend in MOPITT near-surface CO over western Maryland.

Surface concentrations of CO show downward trends over the United States driven by emission reductions (EPA, 2011), consistent with the space-based trends. Decreases in the partial column of CO from FTIR stations in Europe also show decreases from 1996 to 2006, consistent with emissions decreases (Angelbratt et al., 2011). Yoon...
and Pozzer (2014) found that a model simulation of 2001 to 2010 reproduced negative
trends in surface CO over the eastern U.S. and western Europe, but showed a positive
trend in surface CO over southern Asia.

The cause of the negative trend over China seen in MOPITT and AIRS data is
uncertain. The trend is consistent with the results of Li and Liu (2011), who found
decreases in surface CO measurements in Beijing, and with decreases in CO emissions in
2008 inferred from the correlation of CO with CO₂ measured at Hateruma Island
(Tohjima et al., 2014) and at a rural site in China (Wang et al., 2010). Yumimoto et al.
(2014) used inverse modeling of MOPITT data to infer a decrease in CO emissions from
China after 2007. The 2008 Olympic Games and the 2009 global economic slowdown
led to reductions in CO (Li and Liu, 2011; Worden et al., 2012). However, the negative
trend in MOPITT CO is inconsistent with the rising CO emissions of the MACCity and
REAS inventories. Inverse modeling of MOPITT version 6 data yields a negative trend
in CO emissions from China and a larger global decline in CO emissions than that found
in the MACCity inventory (Yin et al., 2015).

This study examines whether global hindcast simulations can reproduce the trends
and variability in carbon monoxide seen in the MOPITT record. We examine the role of
averaging kernels and the contribution of trends at different altitudes to the trends
observed by MOPITT. We then examine the impact of OH variability on the simulated
trends in CO.

2. Methods

2.1. MOPITT

The MOPITT instrument onboard the Terra Satellite provides the longest satellite-
based record of atmospheric CO, with observations available from March 2000 to
present. It provides nearly global coverage every three days (Edwards et al., 2004). We
use the monthly Level 3 daytime column data from the Version 5 TIR product, which has
negligible drift in the bias over time (Deeter et al., 2013). The level 3 data is a gridded
product and includes the a priori and averaging kernel for each grid box. Supplemental
Figure S1 shows the MOPITT column averaging kernels averaged over four regions. The
column averaging kernels depend on the observed scene, and vary year to year as well as
seasonally. The dependence of the column averaging kernels on the CO mixing ratio profile (Deeter, 2009) explains the high values in the lower troposphere over eastern China in winter.

We calculate trends and de-seasonalized anomalies for the Eastern U.S., Europe, and eastern China regions described by Worden et al. (2013). Trends that differ from zero by more than the two-sigma uncertainty on the trend are considered statistically significant. We account for autocorrelation of the data for a one-month lag when calculating the uncertainty on the trends. We calculate the annual cycle by fitting the data with a series of sines and cosines as well as the linear trend, and then remove the annual cycle to obtain the de-seasonalized anomalies. Months with no MOPITT data or only a few days of MOPITT data are excluded from the trend analysis. This includes May-August of 2001 and August-September of 2009. We report the MOPITT trends for 2000-2010 for comparison with model simulations, and for 2000-2014 to give a longer-term view of the observed trends.

2.2. Model Simulations

We use a suite of chemistry climate model (CCM) and chemical transport model (CTM) simulations to interpret the observed trends. The Global Modeling Initiative (GMI) CTM includes both tropospheric (Duncan et al., 2007) and stratospheric (Strahan et al., 2007) chemistry, including over 400 reactions and 124 chemical species. Meteorology for the GMI simulations comes from the Modern-Era Retrospective Analysis for Research and Applications (MERRA) (Rienecker et al., 2011). The GEOS-5 Chemistry Climate Model (GEOSCCM)(Oman et al., 2011) incorporates the GMI chemical mechanism into the GEOS-5 atmospheric general circulation model (AGCM). The GEOSCCM simulations are forced by observed sea surface temperatures (SSTs) from Reynolds et al., 2002).

The Community Earth System Model, CESM1 CAM4-chem, includes 191 chemical tracers and over 400 reactions for both troposphere and stratosphere (Tilmes et al., 2016). The model can be run fully coupled to a free-running ocean, with prescribed SSTs, or with nudged meteorology from GEOS-5 or MERRA analysis. CESM1 CAM4-chem is

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further coupled to the land model, providing biogenic emissions from the Model of Emissions and Aerosols from Nature (MEGAN), version 2.1 (Guenther et al., 2012).

Several simulations were conducted as part of the Chemistry-Climate Model Initiative (CCMI) project (Eyring et al., 2013). These include the Ref-C1 simulation of the GEOSCCM and a Ref-C1 CESM1 CAM4-Chem simulation, hereafter called G-Ref-C1 and C-Ref-C1, respectively, and the Ref-C1-SD simulation of the GMI CTM. Both the Ref-C1 and the Ref-C1-SD simulations use time-dependent anthropogenic and biomass burning emissions from the MACCity inventory (Granier et al., 2011), but the Ref-C1-SD simulations use specified meteorology while the Ref-C1 simulations run with prescribed SSTs. The MACCity inventory linearly interpolates the decadal anthropogenic emissions from the ACCMIP inventory (Lamarque et al., 2010) for 2000, and the RCP8.5 emissions for 2005 and 2010, to each year in between. The MACCity biomass burning emissions have year-to-year variability based on the GFED-v2 (van der Werf et al., 2006) inventory. From 2000 to 2010, CO emissions in the MACCity inventory decreased from 31 to 11 Tg yr\(^{-1}\) over the eastern U.S., from 97 to 59 Tg yr\(^{-1}\) over Europe, and increased from 56 Tg to 72 Tg yr\(^{-1}\) over eastern China.

Given the uncertainty in CO emissions, we conduct a GMI CTM simulation using an alternative time-dependent emissions scenario, called AltEmis. This simulation is described in detail in (Strode et al., 2015b). Briefly, anthropogenic emissions include time-dependence based on EPA (http://www.epa.gov/ttn/chief/trends/index.html), the REAS inventory (Ohara et al., 2007), and EMEP (http://www.ceip.at/ms/ceip_home1/ceip_home/webdab/emepdatabase/reported_emissionndata/), and annual scalings from van Donkelaar et al., (2008). Biomass burning emissions are based on the GFED3 inventory (van der Werf et al., 2010). While the regional emission trends in this simulation are of the same sign as in the Ref-C1 case, the magnitude of the negative trends over the U.S. and Europe are smaller and the positive trend over China is larger, leading to a positive global trend (Fig. 1). We also conduct a sensitivity study called EmFix with anthropogenic and biomass burning emissions held constant at year 2000 levels. Table 1 summarizes the simulations used in this study.

We regrid the model output to the MOPITT grid and convolve the simulated CO with the MOPITT averaging kernels and a priori in order to compare the simulated and
observed CO columns. The averaging kernels are space and time dependent. We use the following equation from Deeter et al. (2013):

\[ C_{\text{sim}} = C_0 + a (\chi_{\text{mod}} - \chi_0) \]  

(1)

where \( C_{\text{sim}} \) and \( C_0 \) are the simulated and a priori CO total columns, respectively, \( a \) is the total column averaging kernel, and \( \chi_{\text{mod}} \) and \( \chi_0 \) are the modeled and a priori CO profiles, respectively. The column averaging kernel is calculated from the standard averaging kernel matrix, which is based on the log of the CO concentration profile, following the method of Deeter (2009):

\[ a_j = \frac{K}{\log_{10}e} \sum \Delta p_i v_{\text{RTV},i} A_{ij} \]  

(2)

where \( \Delta p_i \) and \( v_{\text{RTV},i} \) are the pressure thickness and retrieved CO concentration, respectively, of level \( i \), \( A \) is the standard averaging kernel matrix, and \( K = 2.12 \times 10^{13} \) molec cm\(^{-2}\) hPa\(^{-1}\) ppb\(^{-1}\).

We deseasonalize the simulated CO columns and calculate their linear trend following the same procedure that we applied to the MOPITT CO. Months that do not have MOPITT data (June-July 2001 and August-September 2009) are excluded from the analysis of the model trends as well.

The Ref-C1 and Ref-C1-SD simulations requested by CCMI extend until 2010. However, the MACCity biomass burning emissions extend only until 2008. CAM4-Chem therefore repeated the biomass burning emissions for 2008 for years 2009-2010. In contrast, the GEOSCCM Ref-C1 and GMI Ref-C1-SD simulations used emissions from GFED3 (van der Werf et al., 2010) for years after 2008. Some simulations were available through 2011, while others ended in 2010. We therefore report results for 2000-2010, but note that extending the analysis through 2011 does not alter the conclusions.

3. Results

3.1. Trends over Europe, the United States, and the Northern Hemisphere

The hindcast simulations driven by MACCity emissions (G-Ref-C1, Ref-C1-SD, and C-Ref-C1) show negative trends in CO over the U.S. and Europe that agree with the observed slope from MOPITT within the uncertainty (Fig. 2, Table 2). The MOPITT
trends for both regions are statistically significant for both regions, as shown by Worden et al. (2013). These results are consistent with the findings of Yin et al. (2015), whose inversion of MOPITT data showed a posteriori trends in CO emissions over the U.S. and western Europe that were consistent with but slightly larger than the a priori trends. The EmFix hindcast shows a positive, though non-significant, trend for both regions, indicating that the decrease in CO emissions is necessary for reproducing the downward trend in the CO column. The AltEmis simulation fails to produce the negative trends, despite including negative trends in regional emissions for both the U.S. and Europe. The impact of these negative regional trends is insufficient to overcome the positive global emission trend in the AltEmis scenario (Fig. 1), leading to positive trends in CO.

Figure 2 also reveals a negative bias in the simulated CO column between the models and MOPITT. A low bias in simulated CO at northern latitudes is often present in global models (Naik et al., 2013), and may indicate a high bias in northern hemisphere OH (Strode et al., 2015a) or CO dry deposition (Stein et al., 2014), as well as an underestimate of CO emissions.

The deseasonalized anomalies in the MOPITT and simulated CO columns are shown in Fig. 2b,d, and the correlation coefficient between the observed and simulated monthly anomalies are presented in Table 2b. The highest correlations are for the AltEmis and Ref-C1-SD simulations of the GMI CTM. This result is consistent with the use of year-specific meteorology, which we expect to better match the transport of particular years. The lowest correlations are for the EmFix simulation. This is expected since the EmFix simulation does not include inter-annual variability (IAV) in biomass burning. The IAV in biomass burning makes a large contribution to the IAV of CO (Voulgarakis et al., 2015).

The role of biomass burning in driving the CO variability is even more evident at the hemispheric scale. Figure 2g,h shows the anomalies in MOPITT and the simulations for the northern hemisphere (0-60N). The EmFix simulation shows almost no correlation, while the other simulations have correlation coefficients exceeding 0.6 (Table 2). The role of changing anthropogenic emissions is also evident, as the Ref-C1-SD simulation captures the 2008-2009 dip in the CO column while the EmFix simulation does not. Gratz et al. (2015) found decreasing CO concentrations at Mount Bachelor Observatory
in Oregon during spring for 2004-2013, which they attribute to reductions in emissions leading to a lower hemispheric background. We also note that Ref-C1-SD and G-Ref-C1 have similar correlations with the observed variability for the northern hemisphere (Table 2), indicating that transport differences are less important for variability at the hemispheric scale.

### 3.2. Trend over China

Observations from MOPITT show a negative trend in the CO column over eastern China for 2002-2012 (Worden et al., 2013). The negative trend for the years 2000-2014 exceeds that for 2000-2010 (Table 2), showing that it is not driven solely by temporary emission reductions in 2008. Our simulations do not reproduce this trend, and instead show increases in the CO column (Fig. 2e), which is expected given that CO emissions from China increase in four of the five simulations. The anomalies (Fig. 2f) show that the discrepancy in the simulated versus observed trends is driven largely by the failure of the simulations to capture the 2008 dip in the CO column, leading to an overestimate that continues through 2010. This suggests emission reductions in China during this time period are not adequately captured by the emission inventories. However, the good agreement between the observed and simulated decreases in CO for the northern hemisphere as a whole (Fig. 2g,h) suggest that on a global scale, the emission time series is reasonable. Consequently, we examine several other factors that may contribute to the difference in sign between the MOPITT and simulated CO trends.

Regional trends in CO are expected to vary with altitude, with surface concentrations most heavily influenced by local emissions. MOPITT TIR retrievals have higher sensitivity to CO in the mid-troposphere than at the surface (Deeter et al., 2004), so the trend in the MOPITT CO column will be weighted towards the trends in free tropospheric CO rather than near-surface CO. We quantify this impact on our Ref-C1-SD CO column trends by comparing the trend in the pure-model CO column with that of the simulated column convolved with the MOPITT averaging kernels.

The simulated CO trend over eastern China for 2000-2010 is positive (but not significant) both with and without the averaging kernels, but application of the MOPITT kernels increases the positive trend from $1.3 \times 10^{16}$ mole cm$^{-2}$ yr$^{-1}$ to $1.4 \times 10^{16}$ mole cm$^{-2}$ yr$^{-1}$. This result is initially surprising since we expect trends in the mid-troposphere to be
more strongly influenced by the decrease in the hemispheric CO background. Indeed, the
trends in CO concentration over eastern China simulated in Ref-C1-SD switch from
positive in the lower troposphere to negative in the middle and upper troposphere.
However, the application of the kernels results in more positive (or less negative) trends
in all regions.

Yoon et al. (2013) show that since the averaging kernels vary over time, a bias
between the true atmosphere and the a priori assumed by MOPITT can lead to an
artificial trend in the retrieved CO. Similarly, the bias between the average simulated CO
concentrations and the MOPITT a priori, evident in Figure 2, can lead to an artifact in the
simulated CO trend when the simulation is convolved with the MOPITT averaging
kernels. This is due to the changing contribution of the a priori when the vertical
sensitivity (averaging kernel) is varying in time. MOPITT vertical sensitivity varies with
time due to instrument degradation as well as the change in CO abundance. The bias in
CO varies with altitude, so if the vertical sensitivity described by the averaging kernel
changes, this will change the value of the convolved CO column even if there were no
changes in the CO profile. Furthermore, changes in the averaging kernel result in more
or less weight placed on the a priori versus the CO simulated by the model. Thus, a
difference between the a priori and the model means that placing more (or less) weight on
the a priori will change the resulting value of $C_{\text{sim}}$. Since the a priori profiles and
columns are constant in time, taking the time derivative of equation 1 yields:

$$\partial C_{\text{sim}} / \partial t = a \left( \partial x_{\text{mod}} / \partial t \right) + \partial a / \partial t \left( x_{\text{mod}} - x_0 \right) \quad (3)$$

The second term on the right hand side shows that the larger the bias between the
modeled CO and the a priori, the larger the impact of the changing averaging kernel.

We quantify this effect by convolving the simulated CO for each year with the
MOPITT averaging kernels for the year 2008, thus removing the effect of the time-
dependence of the averaging kernels. The resulting trend, $0.56 \times 10^{16} \text{ molec cm}^{-2} \text{ yr}^{-1}$, is
less positive than the pure model trend or the original simulated trend. Thus, accounting
for the time-dependence of the averaging kernels convolved with model bias reduces but
does not eliminate the discrepancy with the observed trend. Other regions also show a
more negative trend when the same averaging kernel is applied to the model results for
all years. The large bias in CO at middle and high northern latitudes commonly seen in
modeling studies thus impacts the ability of models to reproduce and attribute observed
trends in satellite data.

Figure 2 and Table 2 also show a positive trend in the GMI EmFix simulation for
eastern China. This larger trend in the EmFix simulation than the Ref-C1-SD simulation
dicates that the net decrease in emissions contributes to decreasing CO over eastern
China, consistent with the observed negative trend, but other factors in the model cause
an increase in CO over eastern China even when all emissions are constant. The trend in
the EmFix simulation thus contributes to the erroneous sign of the trend in the GMI
simulations. The trends in the EmFix simulation for the northern hemisphere average and
the eastern U.S. and Europe are positive as well (Table 2). We examine their cause in the
next section.

3.3. Contribution of OH Interannual Variability

Since the EmFix simulation shows a positive trend in the northern hemisphere, we
next examine the variability in the CO sink, OH. We also examine variability in the total
ozone column, since overhead ozone is a major driver of OH variability (Duncan and
Logan, 2008). Figure 3 shows the variability in CO and OH in the EmFix simulation.
The positive and negative anomalies in CO correspond with the negative and positive
anomalies, respectively, in OH. The anomalies in OH are in turn inversely related to
anomalies in the total ozone column. The correlation coefficient between OH and
column ozone is -0.53 for the 15°S-15°N average, -0.72 for the 15°-25°N average, and -
0.75 for the 30°-60°N average. The large NH ozone anomaly in 2010, in particular, leads
to a large anomaly in OH and thus CO. This OH anomaly extends from the northern
tropics to the midlatitudes. The large CO anomaly near the end of the time series
contributes to the apparent 11-year trend. We note that since the lifetime of CO is several
months, CO anomalies are not expected to have a one-to-one correspondence with the
OH anomalies.

The large anomaly in the simulated total ozone column in 2010 is overestimated
compared to observations. Figure 4 shows the time-dependence of the total ozone
column from 30°-60°N in EmFix compared to SBUV data (Frith et al., 2014). While the
observations show an anomaly in 2010, the magnitude is smaller than that produced by
Steinbrecht et al. (2011) attribute the 2010 anomaly in northern midlatitude ozone observations to a combination of an unusually strong negative Arctic Oscillation and North Atlantic Oscillation and the easterly phase of the quasi-biennial oscillation.

While the impact of OH interannual variability on the apparent trend in CO is clear in the EmFix simulation, this source of variability is partially masked by large interannual variability in CO emissions in the other simulations. We examine the correlation between the de-trended and deseasonalized CO anomalies from 10°S-10°N in the Ref-C1-SD simulation and the CO emissions as well as the simulated OH and column ozone. Since the CO emitted in a given month can influence concentrations for several subsequent months, we use a 3-month smoothing of the emission time series. We find a high correlation (r=0.88) between the CO anomalies and the CO emissions. This correlation is also evident in the MOPITT data, as the MOPITT CO anomalies have a correlation of r=0.70 with the emissions. Figure 5 shows the strong relationship between the simulated CO anomalies and the CO emissions. However, the colors in Fig. 5 indicate that the scatter for a given level of emissions is often linked to the OH anomalies, with low/high OH anomalies leading to CO that is higher/lower than would be predicted just from the CO emissions. We find that the 10°S-10°N OH in the Ref-C1-SD simulation is anticorrelated with CO (r=−0.62) and with the total ozone column (r=−0.68). Consequently, the simulated ozone column plays a role in modulating tropical CO variability even when variable CO emissions are included, although the emissions still play the strongest role.

4. Conclusions

We conducted a series of multi-year simulations to analyze the causes of the negative trends in MOPITT CO reported by Worden et al. (2013). Both CTM and CCM simulations driven by the MACCity emissions reproduce the observed trends over the eastern U.S. and Europe, providing confidence in the regional emission trends.

None of the simulations reproduce the observed negative trend over eastern China. This negative trend persists even with the MOPITT data extended out to 2014. The MOPITT averaging kernels are weighted towards the free troposphere, where the relative
importance of hemispheric versus local trends is greater. However, our simulations indicate that this effect is insufficient to explain the negative trends over China. While this likely indicates a too positive emission trend for China, several other factors play a role in the model-observation mismatch. We find that the time-dependent MOPITT averaging kernels, combined with the low bias in simulated CO, provides a positive component to the simulated trends. Large anomalies in the simulated ozone column in the GMI CTM simulations also contribute a positive component to the northern hemisphere trends due to their impact on OH.

Variability in emissions is the primary driver of year-to-year variability in simulated CO, but OH variability also plays a role. The simulated OH is anti-correlated with both CO and the total ozone column, highlighting the importance of realistic overhead ozone columns for accurately simulating CO variability and trends. In addition, further work is needed to understand recent changes in CO emissions from China.

Acknowledgements

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<table>
<thead>
<tr>
<th>Simulation</th>
<th>Model</th>
<th>Meteorology</th>
<th>Anthropogenic Emissions</th>
<th>Biomass Burning Emissions</th>
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<tr>
<td>G-Ref-C1</td>
<td>GEOSCCM</td>
<td>internally derived</td>
<td>MACCity, MACCity,</td>
<td>GFED3 (2009-2010)</td>
</tr>
<tr>
<td>C-Ref-C1</td>
<td>CAM4-Chem</td>
<td>internally derived</td>
<td>MACCity, MACCity,</td>
<td>then repeat 2008</td>
</tr>
<tr>
<td>Ref-C1-SD</td>
<td>GMI</td>
<td>MERRA</td>
<td>MACCity</td>
<td>Same as GEOSCCM</td>
</tr>
<tr>
<td>EmFix</td>
<td>GMI</td>
<td>MERRA</td>
<td>Fixed at 2000</td>
<td>Fixed at 2000</td>
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<tr>
<td>AltEmis</td>
<td>GMI</td>
<td>MERRA</td>
<td>Strode et al [2015]</td>
<td>GFED3</td>
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### Table 2: Regional Trends and Correlations

#### a. Trends\(^1,2\)

<table>
<thead>
<tr>
<th>Years</th>
<th>E. USA</th>
<th>Europe</th>
<th>E. China</th>
<th>N. Hemisphere</th>
</tr>
</thead>
<tbody>
<tr>
<td>C-Ref-C1(^3)</td>
<td>2000-2010</td>
<td>-3.4 (0.54)</td>
<td>-2.9 (0.50)</td>
<td>1.4 (1.4)</td>
</tr>
<tr>
<td>Ref-C1-SD(^3)</td>
<td>2000-2010</td>
<td>-2.4 (0.53)</td>
<td>-1.6 (0.59)</td>
<td>1.4 (1.1)</td>
</tr>
<tr>
<td>EmFix(^3)</td>
<td>2000-2010</td>
<td>1.3 (0.55)</td>
<td>1.5 (0.44)</td>
<td>2.1 (0.87)</td>
</tr>
<tr>
<td>AltEmis(^3)</td>
<td>2000-2010</td>
<td>0.71 (0.73)</td>
<td>0.74 (0.66)</td>
<td>3.8 (1.4)</td>
</tr>
</tbody>
</table>

10\(^{16}\) molec cm\(^{-2}\) yr\(^{-1}\)

1-sigma uncertainty given in parentheses

\(^3\)Simulation results convolved with MOPITT averaging kernel and a priori

#### b. Correlation coefficient (r) with monthly MOPITT anomalies\(^1,2\)

<table>
<thead>
<tr>
<th>Years</th>
<th>E. USA</th>
<th>Europe</th>
<th>E. China</th>
<th>N. Hemisphere</th>
</tr>
</thead>
<tbody>
<tr>
<td>G-Ref-C1</td>
<td>2000-2010</td>
<td>\textbf{0.26}</td>
<td>\textbf{0.39}</td>
<td>\textbf{0.061}</td>
</tr>
<tr>
<td>C-Ref-C1</td>
<td>2000-2010</td>
<td>0.23</td>
<td>\textbf{0.36}</td>
<td>0.18</td>
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<tr>
<td>Ref-C1-SD</td>
<td>2000-2010</td>
<td>\textbf{0.43}</td>
<td>\textbf{0.51}</td>
<td>\textbf{0.39}</td>
</tr>
<tr>
<td>EmFix</td>
<td>2000-2010</td>
<td>0.10</td>
<td>0.21</td>
<td>0.071</td>
</tr>
<tr>
<td>AltEmis</td>
<td>2000-2010</td>
<td>\textbf{0.55}</td>
<td>\textbf{0.59}</td>
<td>\textbf{0.48}</td>
</tr>
</tbody>
</table>

Correlations are calculated from the de-trended and de-seasonalized time series.

Statistically significant correlations at the 95% confidence level are indicated in bold.
<p>| | | | | | |</p>
<table>
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<tr>
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</tr>
</thead>
<tbody>
<tr>
<td>G-Ref-C1</td>
<td>2000-2011</td>
<td>-2.4 (0.33)</td>
<td>-1.9 (0.36)</td>
<td>1.9 (0.97)</td>
<td>-0.90 (2.5)</td>
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<tr>
<td>Ref-C1-SD</td>
<td>2000-2011</td>
<td>-2.1 (0.43)</td>
<td>-1.7 (0.51)</td>
<td>1.3 (1.0)</td>
<td>-0.89 (2.6)</td>
</tr>
<tr>
<td>EmFix</td>
<td>2000-2011</td>
<td>1.3 (0.43)</td>
<td>1.3 (0.39)</td>
<td>2.0 (0.79)</td>
<td>0.91 (2.1)</td>
</tr>
<tr>
<td>AltEmis</td>
<td>2000-2011</td>
<td>0.56 (0.59)</td>
<td>0.50 (0.58)</td>
<td>3.3 (1.3)</td>
<td>0.89 (2.9)</td>
</tr>
</tbody>
</table>

| MOPITT   | 2000-2011     | -2.5 (0.55)| -1.9 (0.59)| -2.9 (1.5)| -1.5 (2.4) |

|          |               |            |            |            |            |
| G-Ref-C1 | 2000-2011     | 0.25       | 0.37       | 0.046      | 0.72       |
| Ref-C1-SD| 2000-2011    | 0.41       | 0.51       | 0.36       | 0.73       |
| EmFix    | 2000-2011     | 0.095      | 0.24       | 0.080      | 0.094      |
| AltEmis  | 2000-2011     | 0.53       | 0.59       | 0.47       | 0.69       |
Figure 1: Trends in the CO emissions used in the Ref-C1 and Ref-C1-SD simulations (blue bars) and AltEmis simulation (purple bars) over 2000-2010 for the United States, Europe, China, and the world.
Figure 2: The time series and trends (left column) and de-seasonalized monthly anomalies (right column) of the CO column from MOPITT (black), the MOPITT a priori (gray), and simulated by G-Ref-C1 (red), Ref-C1-SD (blue), EmFix (green), C-Ref-C1 (orange), and AltEmis (purple) for 2000-2010. The regions shown are (a) Europe (0°-60°N) and (b) Europe anomaly, (c) East USA, (d) East USA anomaly, (e) East China, (f) East China anomaly, (g) N Hemisphere (0°-60°N), and (h) N Hemisphere anomaly.
15°E, 45°-55°N), (c,d) eastern U.S.A. (95°-75°W, 35°-40°N), (e,f) eastern China (110°-123°E, 30°-40°N), and (g,h) the northern hemisphere (0°-60°N).

Figure 3: Deseasonalized monthly anomalies in the total ozone column (left), mean tropospheric OH (center), and CO column (right) from the EmFix simulation as a function of latitude and month.
Figure 4: Monthly ozone column (a) and de-seasonalized ozone column anomaly (b) in SBUV data (black) and the EmFix simulation (green) for 30°-60°N.

Figure 5: Monthly simulated CO column anomalies from the Ref-C1-SD simulation as a function of CO emissions for 10°S-10°N. Colors indicate the simulated OH column anomaly for the given month.
Supplemental Figure S1: Regionally averaged monthly column averaging kernels for January (circles) and July (triangles) for 2006 (black), 2008 (red), and 2010 (blue) for (a) the eastern United States, (b) Europe, (c) eastern China, and (d) the northern hemisphere. The region definitions are the same as in Fig. 2.